

# Savings and Subsidies, Separately and Together: Decomposing Effects of a Bundled Anti-Poverty Program\*

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## Abstract

Many anti-poverty programs are “bundled,” in that they have multiple active program components. We shed light on the interplay between two important types of programs by randomly assigning rural households in Mozambique to subsidies for modern agricultural inputs, formal savings facilitation programs (either a “basic” or a “matched” savings program), or both subsidy and savings programs. We take household consumption per capita as the summary welfare measure of interest. Impacts on consumption are very similar, from an economic and statistical standpoint, for all treatment combinations (subsidy alone, either savings program alone, or combinations of subsidies and savings), amounting to 7.7 percent on average. Theoretically, such a pattern can emerge if access to input subsidies affects the uses to which formal savings are put: savings may be used for mainly self-insurance (buffer stocks) when subsidies are available, but otherwise may serve both investment and insurance purposes. Auxiliary predictions of the theory are borne out in analyses of treatment effects on consumption variance and on the responsiveness of consumption to shocks. A key gain from combining subsidy and savings interventions appears to be an improved ability to cope with risks that arise from exploitation of high-return economic opportunities.

savings, subsidies, risk, agriculture, Mozambique

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# 1 Introduction

While there is a continually growing body of evidence on the impacts of development programs implemented on their own, there is comparatively little evidence on the complementarities between multiple interventions implemented simultaneously. It is important to identify such complementarities, because interventions nearly always occur alongside other concurrent programs. In addition, major development proposals often involve a large number of concurrent interventions. For example, Sachs (2005) proposes multiple simultaneous interventions in each beneficiary country, and justifies this in part on the basis of positive complementarities across interventions. Banerjee et al. (2015) show positive, persistent impacts of a multifaceted development program for the “ultrapoor” across six countries involving resource transfers, skills training, savings, and health education. At this stage of research on anti-poverty programs in developing countries, there is an emerging need for a better understanding of the interplay among the components of bundled interventions.

This paper seeks to shed light on the interaction between two types of prominent anti-poverty interventions: agricultural input subsidies and formal savings programs. Large-scale subsidization of modern agricultural inputs (fertilizer and hybrid seeds) has emerged as perhaps the most significant recent development in agricultural policy in Sub-Saharan Africa. Ten countries have implemented input subsidy programs (known as ISPs) in recent decades. In 2011, total expenditures totaled \$1.05 billion, or 28.6% of public agricultural spending in these countries (Jayne and Rashid (2013).) These programs receive substantial budgetary support from international development agencies such as the World Bank. Support for ISPs represents an about-face for many development agencies, which for decades opposed subsidies of this sort (Morris et al. (2007)). Randomized experimental studies of the impact of ISPs (including Duflo et al. (2011), Carter et al. (2014), and Harou et al. (2014)) have found mixed evidence on their effectiveness at raising input utilization, in particular on the question of whether the impacts of temporary subsidies extend over time, beyond the subsidized agricultural season.

There has also been a recent flourishing of empirical evidence on the impacts of facilitating formal savings in developing countries. Savings, in theory, can facilitate accumulation of investment capital as well as buffer stocks that help cope with risk (Kimball (1990), Deaton (1990), Deaton (1991), Deaton (1992), Collins et al. (2009)). Savings programs often provide formal savings facilities to the poor, to complement informal savings. Demirguc-Kunt and Klapper (2013) document that formal savings is strongly positively associated with income, in cross-country comparisons as well as across households within countries. Randomized savings-facilitation interventions have been shown to affect household expenditure composition (Prina (2015)) and to improve asset accumulation (Dupas and Robinson (2013a) and our companion paper Carter et al. (2015b)), the ability to cope with health shocks (Dupas and Robinson (2013b)), and household consumption levels (Brune et al. (forthcoming)).

We conducted a randomized controlled trial in Mozambique that tested for complementarities between input subsidies and a savings-oriented financial services intervention. Study participants were randomly assigned to being offered either a subsidy for modern agricultural inputs, a savings facilitation program, or both. We tested two types of savings facilitation programs: a “basic savings” program (providing only information on savings and access to bank accounts) and a “matched savings” program that in addition incentivized savings with generous matching funds.<sup>1</sup> The research design allows us to estimate the impact

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<sup>1</sup>The matched savings treatment could be thought of as a behavioral “nudge” to initiate formal savings, which might then generate persistence in saving (for example, by facilitating learning-by-doing about the benefits of savings). Previous studies of matched savings programs (often called individual development accounts, or IDAs, in the US) include Boshara (2005), Schreiner and Sherraden (2007), Sherraden and McBride (2010), Sherraden (1988), Sherraden (1991), Grinstein-Weiss et al. (2013b), and Grinstein-Weiss et al. (2013a). Schaner (2015) finds persistent impacts of a randomized matched-savings intervention in Kenya.

of each type of program separately, and also measure complementarities when participants are offered them together.

By “complementary” we mean a situation where the impact of both interventions implemented together is larger than the sum of impacts of these interventions when implemented separately. More generally, we distinguish among three general cases of complementarity, from the standpoint of influencing a widely-used summary measure of household well-being: log consumption per capita. Consider two interventions whose impacts on log consumption when offered separately are, respectively,  $\alpha$  and  $\beta$ , and whose impact when offered together is  $\alpha + \beta + \gamma$ . The parameter  $\gamma$  measures the degree of complementarity. The interventions are *complementary* when the impact of the joint intervention is greater than the sum of impacts when offered separately:  $\gamma > 0$ . If the joint impact is simply equal to the sum of the separate impacts ( $\gamma = 0$ ), we say the interventions are *additive*. The third case is where the joint impact is lower than the sum of the separate impacts ( $\gamma < 0$ ) in which case the interventions are *substitutes*.

We examine impacts of subsidies and savings, separately and together, on household consumption per capita in periods following the treatments. Our first key finding is that from standpoint of raising consumption, subsidies and savings appear to be substitutes, rather than complements. Among respondents not exposed to any savings treatment, subsidies have positive impacts on fertilizer use and ultimately on household consumption levels in subsequent years.<sup>2</sup> The savings treatments themselves (absent the subsidies) also have positive impacts. Strikingly, the impact of the subsidies is no larger when combined with one of the savings programs. The treatment effect, averaged across treatment combinations (subsidies alone, either basic or matched savings alone, or subsidies combined with either savings program), amounts to an increase in household consumption per capita of 7.7 percent, and we cannot reject at conventional levels of statistical significance that the treatments all have equal impacts. Study participants use either treatment to increase consumption to similar extents, but experience no further increases when they get both treatments compared to when they get just one of them.

The absence of complementarity (full substitutability) between subsidy and savings interventions can arise, theoretically, if savings can serve multiple purposes, facilitating asset accumulation and investment, as well as improving the ability to cope with risk. We write down a simple model in which households are potentially liquidity-constrained (hindering their investment) and information-constrained (limiting their use of modern inputs and of formal savings facilities). Introduction of subsidies by themselves can raise consumption by raising agricultural input usage. Introduction of formal savings alone can also raise consumption by helping households accumulate resources for future investment and input use. When offered in tandem, subsidies and savings need not lead to further input use (and consumption gains), but households may use savings to improve consumption smoothing via buffer stock accumulation in formal savings accounts, which leads to utility gains for risk-averse households.

This simple theory makes auxiliary predictions which we test in additional empirical analyses. First, we show that the subsidy treatment, by itself, increases risk, significantly raising the *variance* of consumption, even as it raises the *level* of consumption. Households receiving both the subsidy and the basic savings program experience similar increases in mean returns, but have significantly lower consumption variance. Figure 1 shows this latter result graphically, presenting probability density functions of post-treatment log

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See also Ambler et al. (2015) and Karlan and List (2007) on the impacts of provision of matching funds in different contexts. Research on matching programs and tax credits for saving is also related. Duflo et al. (2006) find positive effects of savings matching programs on savings (also see Bernheim (2003), Choi et al. (2011), Engelhardt and Kumar (2007), Engen et al. (1996), Even and MacPherson (2005), Gale et al. (2005), Huberman et al. (2007), and Papke and Poterba (1995).)

<sup>2</sup>We explore the impact of subsidies absent the savings treatments more thoroughly in a companion paper, Carter et al. (2014).

consumption for the three subsidy treatment groups. PDFs for the subsidy treatments combined with either the basic savings (long dashed line) or matched savings (short dashed line) treatments are visibly more concentrated around their respective means than the PDF for the subsidy-only treatment (solid line). These differences in the variance of consumption are consistent with savings serving as a risk-coping mechanism. More direct evidence of the risk-coping role of savings comes in analysis of the responsiveness of household consumption to agricultural shocks. We find that when households experience negative agricultural shocks, their consumption declines more in the subsidy-only treatment group, compared to any of the treatment groups that involve savings.<sup>3</sup>

In addition, we show that all treatments raise investment in agricultural activities or inputs. The subsidy-only, basic-savings-plus-subsidy, and both matched savings treatments (without and with subsidy) have positive impacts on use of modern fertilizer which statistically significant are roughly similar to one another in magnitude. The exception is the basic-savings-only treatment, which has an essentially zero impact on fertilizer use. This treatment is associated, however, with a statistically significant increase in other agricultural investments and a large (but imprecisely estimated) increase in livestock asset values. It appears, therefore, that consumption gains in that treatment group reflect investments other than fertilizer.

Our work is, of course, related to the vast literature in economics that documents myriad ways in which households in developing countries seek to balance risk and return. When a risk-return tradeoff exists, as is typically the case in agricultural production, households will often seek smoother income at the cost of lowering mean income, by diversifying crops or plot locations, or by making less risky crop and other production choices (Morduch (1993)). The variability of income becomes less of an issue (and households should be more willing to maximize income) if households are able to smooth consumption over time, and there is much evidence that they use a variety of means to do so. They save and dissave (Paxson (1992), Mazzocco (2004)); take out loans (Morduch (1998)); supply more labor (Kochar (1999), Jayachandran (2006)); engage in insurance arrangements, particularly informally within social networks (Townsend (1994), Foster and Rosenzweig (2001), Fafchamps and Lund (2003), Ligon et al. (2002)); receive transfers from migrants (Rosenzweig and Stark (1989), Yang and Choi (2007), Yang (2008), Jack and Suri (2013)); and engage in hybrid credit-cum-insurance arrangements (Udry (1994)). Consumption smoothing is typically far from perfect, however (Fafchamps et al. (1998), Ligon et al. (2002), Kazianga and Udry (2006)), and itself can come at a sacrifice of average income levels, if production assets also serve as buffer stocks (Rosenzweig and Wolpin (1993)).<sup>4</sup> Formal insurance against important sources of income risk can in principle help households make more favorable risk-return trade-offs. There has been particular interest in weather-based index insurance, which pays out on the basis of weather realizations alone and so is immune to adverse selection and moral hazard problems (Carter et al. (2015a)). However, there has been relatively low demand for formal insurance (Gine and Yang (2009), Mobarak and Rosenzweig (2012), Cole et al. (2013), Cai et al. (2015)), though when farmers can be induced to take it up it increases their willingness to take on riskier production activities (Cole et al. (2014), Karlan et al. (2014), Mobarak and Rosenzweig (2014).)

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<sup>3</sup>We also investigate whether reductions in variance may also reflect reductions in inequality due to savings treatments being more likely to benefit poorer households. We find mixed evidence on this front. On the one hand, none of the treatments are particularly “pro-poor”: the benefits from all treatment combinations are larger for households with initially higher consumption per capita, to about the same degree. On the other hand, quantile regression results indicate that the savings treatments (particularly the matched savings treatments) have positive impacts on lower deciles of the consumption distribution than the subsidy-only treatment. Our broad conclusion is that while it appears that at least some of the consumption-variance reduction accompanying the savings treatments is due to risk-reduction, some may also be due to the savings treatments being more pro-poor.

<sup>4</sup>Also related, of course, is the literature on the role of financial markets and public tax-transfer mechanisms in smoothing consumption in developed countries, such as Jappelli and Pistaferri (2011), Asdrubali et al. (1996), and Sorensen and Yosha (1998).

Our results reveal how households seek to balance risk and return in their intertemporal decision-making, suggesting that complementarities between development programs may be in risk management rather than in income maximization. Our results complement those of Cole et al. (2014), Karlan et al. (2014), and Mobarak and Rosenzweig (2014) who find, in a randomized experiments in Ghana and India, that weather-based index insurance makes farmers more willing to take on production risk.<sup>5</sup> To the extent that these and other work also find that index insurance has nontrivial shortcomings (basis risk, imperfect trust, and overweighting of recent rainfall shocks in insurance purchase decisions), our results are useful in showing that a simple program of savings facilitation can also help with household risk-management. Our work also differs from these previous studies in that we are able to characterize the risk-return tradeoffs households are making by examining impacts on the mean and variance of a summary measure of household well-being, per capita consumption.

This paper is organized as follows. Section 2 details the research design, and Section 3 describes the sample, data sources, and basic summary statistics. Section 4 presents the estimates of treatment impacts on consumption and estimates the complementarity of subsidies and savings. In Section 5 we discuss a simple theory that can explain our results. Section 6 provides additional empirical analyses that confirm additional predictions of the theory. Section 7 provides concluding thoughts.

## 2 Research design

We are interested in the impact of agricultural input subsidies, savings facilitation programs, and the interaction of the two. Localities in Manica province were selected to be part of the study on the basis of inclusion in the provincial input voucher program as well as access to a mobile banking program run by Banco Oportunidade de Mocambique (BOM), our partner institution for the savings component of the project. To be accessible to the BOM savings program, which involved scheduled weekly visits of a truck-mounted bank branch, a village had to be within a certain distance of a paved road and within reasonable driving distance of BOM’s regional branch in the city of Chimoio. These restrictions led to inclusion of 94 localities<sup>6</sup> in the larger study, across the districts of Barue, Manica, and Sussundenga.

Within each locality, lists of eligible farmers were created jointly by government agricultural extension officers, local leaders, and agro-input retailers. Individuals were deemed eligible for participation in the study if they met the following criteria: 1) farming between 0.5 hectare and 5 hectares of maize; 2) being a “progressive farmer,” defined as a producer interested in modernization of their production methods and commercial farming; 3) having access to agricultural extension and to input and output markets; and 4) stated interest in the input subsidy voucher (which included paying for the remaining portion of the value of the input package that was not covered by the voucher). Potential study participants were informed that the subsidy voucher would be awarded by lottery to 50% of study participants within each village. Only one person per household was allowed to register as a study participant.

Our study design involves randomization of an agricultural input subsidy voucher at the individual study participant level (within localities), crossed with randomization of savings programs across the 94 localities. Randomization of both the vouchers and the savings programs were conducted by the research team on the

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<sup>5</sup>Vargas-Hill and Viceisza (2012) find similar results in an artefactual field experiment in Ethiopia.

<sup>6</sup>The localities we use were defined by us for the purpose of this project, and do not completely coincide with official administrative areas. We sought to create “natural” groupings of households that had some connection to one another. In most cases our localities are equivalent to villages, but in some cases we grouped adjacent villages together into one locality, or divided large villages into multiple localities.

computer of one of the PIs.

The geographic distribution of localities with respect to the savings treatments is presented in Figure 2. Open circles indicate control (no-savings) localities, open triangles basic savings localities, and filled (black) triangles matched savings localities. The map also indicates the locations of four large towns (Catandica, Manica, Chimoio, and Sussundenga), BOM’s Bancomovil service locations (red stars), and locations of fixed branches (blue stars, all of which are in one of the four towns). BOM’s two fixed branches are located in Chimoio and Manica towns.

## 2.1 Subsidy treatment

The voucher randomization was conducted first. In September through December 2010 (at the beginning of the 2010-2011 agricultural season),<sup>7</sup> vouchers were randomly assigned to 50% of study participants in each locality.

The subsidy voucher randomization was done in the context of a larger nationwide pilot input subsidy program conducted by the Mozambique government. The Manica provincial government agreed to collaborate with our project and allow the randomization of the voucher assignment within the study villages. The voucher provided beneficiary farmers for a subsidy for the purchase of a technology package designed for a half hectare of improved maize production: 12.5 kg of improved seeds (either open-pollinated variety or hybrid) and 100 kg of fertilizer (50 kg of urea and 50 kg of NPK 12-24-12). The market value of this package was MZN 3,163 (about USD 117), of which MZN 2,800 was for the fertilizer component, and MZN 363 was for the improved seed. Farmers were required to co-pay MZN 863 (USD 32), or 27.2% of the total value of the package.<sup>8</sup> In a separate companion paper, Carter et al. (2014), we focus only on the 32 localities randomly selected to be in the “no savings” condition, and therefore did not experience any savings treatment, and analyze impacts of the randomized voucher on the persistence of fertilizer adoption and on household agricultural production. Please refer to that paper for further details on the voucher program and its impacts.

## 2.2 Savings treatments

Later, in April 2011, each of the selected 94 localities was then randomly assigned to either a “no savings” condition or to one of two savings treatment conditions (“basic savings” and “matched savings”), each with 1/3 probability.

### 2.2.1 Basic savings treatment

The first meeting with study participants in the basic savings localities was a financial education session. The training sessions were implemented jointly by the study team and our partner bank, BOM, and covered the benefits of using fertilizer and improved seeds, the importance of saving in order to be able to afford agricultural inputs and other investments, and the use of savings accounts to accumulate buffer stocks for self-insurance. In addition, participants were introduced to BOM and were told how to open and use a savings account.

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<sup>7</sup>The agricultural season in Manica province starts with planting in November and December, with the heaviest rain occurring in December through April, and harvest occurring in May and June. There is a dry period from July through October during which little agricultural activity occurs.

<sup>8</sup>At the time of the study, one US dollar (USD) was worth roughly 27 Mozambican meticals (MZN).

In the first session, participants were asked to form groups of five study participants and select one representative per group. Representatives were offered a t-shirt with the BOM logo and were given the responsibility of maintaining the connection between the bank and the members of their group. Two follow-up sessions, organized between May and July 2011, allowed BOM personnel to check with representatives about the progress of their groups towards opening savings accounts and to address participants' questions and concerns. Representatives were also given more financial education, including additional educational materials to share with their group members at home (a comic and a board game about savings.) At the end of each follow-up session, participants were asked to communicate what they had learned to the rest of their group members. All meetings were organized in the communities, and the representatives were usually offered a meal or a snack during the training. The initial information sessions, to which all participants were invited, and the two follow-ups, which the representatives attended, define the basic savings intervention.

### **2.2.2 Matched savings treatment**

In the matched savings treatment localities, we also implemented all elements of the basic savings treatment described above. In addition, participants were also offered a savings match for savings held at BOM during a defined three-month period in 2011 and 2012.

The matched savings treatment offered a 50% match on the minimum amount that was saved between August 1 and October 31 of 2011 and 2012, with a maximum match of MZN 1500 per individual (approximately USD 56). A flyer was given to savings group representatives, summarizing the rules of the savings match.

The aim of the matched savings treatment was to familiarize the maize farmers with the banking system and encourage them to develop a habit of saving between harvest and planting time, when fertilizer and other inputs are typically purchased. The amount was deposited in beneficiaries' accounts at BOM during the first week of November. The timing of the match program was chosen with the agricultural calendar in mind. A majority of farmers sell most of their maize production before August and purchase their agricultural inputs in November. Although the information sessions emphasized savings to purchase the inputs needed for maize production, once beneficiaries received their the matching funds, they could use the funds for any purpose.

## **3 Sample and data**

### **3.1 Data**

Our sample for analysis in this paper is 1,534 study participants and their households in the 94 study localities. The data used in our analyses come from household survey data we collected over the course of the study.

Surveys of study participants were conducted in person at their homes. Due to uncertainties in the timing of voucher distribution and delays in the creation of the list of study participants at the start of the 2010-2011 agricultural season, it was not feasible to conduct a baseline survey prior to the voucher lottery at the end of the 2010 calendar year. Our first survey was in April 2011, which before the savings treatments but after the voucher treatment. While this is therefore not a true baseline survey with respect to the voucher subsidy treatment, it does include questions on time-invariant variables (e.g., gender) as well as retrospective questions on respondents' pre-voucher-lottery agricultural outcomes and behaviors (relating to the the 2009-2010 season). Only time-invariant variables or outcomes reported retrospectively about the

previous agricultural season will be used as control variables and in the balance tests (Table 2, to be discussed below).

Follow-up surveys were implemented in September 2011, September 2012, and July-August 2013. These follow-up surveys were timed to occur after the May-July annual harvest period, so as to capture fertilizer use, production, and other outcomes related to that harvest. The surveys included modules on savings, consumption, assets, investments, fertilizer use, and agricultural production.

A central outcome variable is daily consumption per capita, which we take as our summary measure of well-being. In each survey round, we calculate the total value (in meticaís) of daily consumption in the household, and divide by the number of household members. Total consumption is the sum of a large number of detailed consumption items, whether purchased or consumed from home production. Detailed consumption items are collected for different time windows, depending on the item: over the past 7 days (food items), 30 days (non-food items such as personal items, transportation, utilities, and fuel), and 12 months (household items, clothing and shoes, health expenditures, ceremonies, education). This is a non-durables consumption measure (durable goods are not included, but are examined separately in the empirical analysis.) Consumption by item is converted to daily frequency before summing to obtain total consumption.

### 3.2 Summary statistics and balance tests

Table 2 presents means (standard deviations in parentheses) of baseline variables for the study households, and tests of balance on these variables across study participants in the control group and treatment groups T1 through T5. Sample household heads are roughly 85% male, and about three-quarters are literate. Given that the sample is composed of farmers considered “progressive” by provincial extension agents, these figures are somewhat higher than Manica province households overall, among which 66% of household heads are male and 45% are literate.<sup>9</sup> During the 2009-2010 season, prior to the study, households farmed between three and four hectares of land, and roughly one-fifth used fertilizer on at least one of their maize fields.

The table also tests balance between treatment and control groups for variables that should not be affected by the subsidy treatment (e.g., education of the household head), or agricultural variables related to the 2009-10 agricultural season (the season prior to our study.) Columns for each of treatment groups T1 through T5 report in brackets the p-values of the F-tests of pairwise equality of the mean in that treatment group and the mean in the control group. Section A presents the balance tests for the baseline variables in levels, and Section B is analogous with a subset of variables (continuous variables related to production) specified in logs.

In Section A, out of 90 such pairwise comparisons, three differences vis-a-vis the control group are statistically significantly different from zero at the 10% level, and four are statistically significantly different from zero at the 5% level. In Section B, out of 50 pairwise comparisons, two are significant at the 10% level and one at the 5% level. This number of statistically significant differences is no larger than would be expected to arise by chance.

Because our outcome variables of interest are obtained from our follow-up surveys, it is important to examine whether attrition from the survey is correlated with treatment (as any such correlation could potentially lead to biased treatment effect estimates.) We examine the relationship between treatment and attrition by regressing an indicator for attrition on treatment indicators and stratification cell fixed effects, and results are in Appendix Table 1. There are 1,589 observations in each regression, representing all the

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<sup>9</sup>The Manica data used for comparison is from the 2007 “Terceiro Recenseamento Geral da População e Habitação,” provided by Mozambique’s National Institute of Statistics, accessible online at [http://www.ine.gov.mz/home\\_page/censo2007](http://www.ine.gov.mz/home_page/censo2007).



individuals included in the lottery for subsidy vouchers at the end of 2010. Surveys of all households of study participants were attempted in each subsequent survey round (in other words, attrition was not cumulative), so all attrition rates reported are vis-à-vis that initial sample. Attrition is 9.9% in the first (2011) follow-up survey, 10.9% in the second (2012) round, and 6.9% in the final (2013) round. Because we combine data from the second and third rounds in our main results tables (and use data from one round when the other is not available), another relevant statistic is that only 3.5% of respondents attrited from both the second and third rounds. There is no evidence of economically or statistically significant differentials in attrition related to treatment. Some coefficients on treatment are somewhat larger for attrition in the second round, with the coefficient the matched savings-only treatment (T4) being relatively large (4.7 percentage points) and significant at the 10% level. But the most important test is in the fourth column, for attrition from both the second and third rounds. In this case none of the coefficients on treatment indicators large or statistically significantly different from zero. Attrition bias is therefore not likely to be a concern in our context.

## 4 Treatment effects and complementarities with respect to the level of consumption

Random assignment to the various treatments allows us to estimate their causal impacts. We are interested in treatment effect estimates as well as estimates of the complementarity between voucher and savings treatments. We take our focal outcome of interest to be per capita consumption in the household, a widespread summary measure of well-being.

For post-treatment outcome  $Y_{ijk}$  for study participant  $i$  in locality  $j$  and stratification cell  $k$ , we estimate the following regression equation:

$$Y_{ijk} = \zeta + \alpha V_{ijk} + \beta_b B_{jk} + \beta_{bv} BV_{ijk} + \beta_m M_{jk} + \beta_{mv} MV_{ijk} + X'_{ijk} \delta + \theta_k + \epsilon_{ijk} \quad (1)$$

$V_{ijk}$ ,  $B_{jk}$ ,  $BV_{ijk}$ ,  $M_{jk}$ , and  $MV_{ijk}$  are indicator variables for assignment to the various treatments, respectively: subsidy-only, basic savings-only, basic savings plus subsidy, matched savings-only, and matched savings plus subsidy.<sup>10</sup> Coefficients on these indicator variables ( $\alpha$ ,  $\beta_b$ ,  $\beta_{bv}$ ,  $\beta_m$ , and  $\beta_{mv}$ ) are key parameters of interest: the ITT estimates of impact of each respective treatment condition.

Linear combinations of the coefficients on treatment conditions provide estimates of the complementarity between each savings treatment and the subsidy. We can decompose the treatment effect of the basic savings plus subsidy treatment into  $\beta_{bv} \equiv \alpha + \beta_b + \gamma_b$ , where  $\gamma_b$  represents the complementarity between the basic savings and subsidy treatments. Analogously, for the matched savings treatment effect the decomposition is  $\beta_{mv} \equiv \alpha + \beta_m + \gamma_m$ , with  $\gamma_m$  representing the complementarity between the matched savings and subsidy treatments.

The complementarity terms are therefore given by linear combinations of coefficients  $\gamma_b = \beta_{bv} - \alpha - \beta_b$  and  $\gamma_m = \beta_{mv} - \alpha - \beta_m$ . Positive complementarity terms ( $\gamma > 0$ ) will be taken to indicate *complementarity* of the treatments: the impact of the joint treatment is larger than the sum of impacts when the treatments are provided separately. Negative complementarity terms ( $\gamma < 0$ ), wherein the impact of the joint treatment is less than the sum of impacts of the separate treatments, will indicate *substitutability* of the treatments. A complementarity term close to zero will indicate that the two treatments are simply *additive*, neither

<sup>10</sup>The “ $i$ ” subscript on the voucher indicator and for the indicators for the voucher interacted with a savings treatment highlights that vouchers were randomized at the individual level, in contrast to the savings treatments which were randomized at the locality level.

complementary nor substitutes. In key regression results tables, we report these complementarity terms (and standard errors), below the corresponding regression.

$\theta_k$  are stratification cell fixed effects representing the groupings of nearby localities within which treatments were randomized (recall that savings treatments were assigned so as to come as close as possible to a uniform distribution of savings treatments within each locality group.)  $X_{ijk}$  is a vector of pre-treatment household-level control variables (all variables in Table 2), which absorb residual variation and help improve precision of the treatment effect estimates. Randomization of the savings treatment is at the locality level, so we report standard errors clustered at the level of the 94 localities (Moulton (1986).)

The outcome variable of interest, consumption per capita, as well as other outcome variables, have substantial noise and relatively low autocorrelation. We follow McKenzie (2012) and estimate treatment effects on the average of post-treatment outcomes across multiple periods, specifically across the 2012 and 2013 follow-up surveys.<sup>11</sup> This helps average out the noise in these variables, and allows greater statistical power.

Dependent variables in the regressions are expressed in logs. This helps moderate the undue influence of extreme values, as well as allowing coefficient estimates to be interpreted (approximately) as percentage changes.<sup>12</sup> Control variables that are not dummy variables are also expressed in logs.

## 4.1 Impacts on consumption and assets

We examine impacts on our primary summary measure of household well-being, per capita daily consumption. For comparison, we also examine a common alternative summary well-being measure, total household assets. Both these variables are specified in log form. We also examine impacts on an index of consumption and assets, since households may certainly value increases in both. Examining this index also helps deal with concerns about improper inference when examining multiple outcome variables. The index of consumption and assets is the average of the log consumption and log assets variables after each has been normalized the mean and standard deviation in the pure control group (as recommended by Kling et al. (2007)).

Table 3 presents regression results from estimation of equation 1. In the consumption regression (column 1), each single-treatment coefficient is similar in magnitude (ranging from 0.084 for the subsidy treatment to 0.099 for the matched savings treatment), and each is statistically significant from zero at the 5% level. The effect of the matched savings plus subsidy treatment is positive and very similar in magnitude (0.088) to the single-treatment effects, and is statistically significantly different from zero at the 5% level. The effect of the basic savings plus subsidy treatment is smaller than the other treatment effects, at 0.037, and is not statistically significantly different from zero at conventional levels. That said, we cannot reject at conventional levels that all the treatment effects are equal (p-value = 0.493). We also cannot reject in any pairwise comparison that the effect of the joint basic savings and subsidy treatment is different from the other treatment effects at conventional statistical significance levels. We do reject the hypothesis that treatment effects are jointly zero (p-value = 0.034).

Complementarity terms are reported below the main regression coefficients. Both complementarity parameters  $\gamma_b$  and  $\gamma_m$  are negative and are at least as large in absolute value as the main effects of each

<sup>11</sup>To maximize sample size, in cases where the value from one year is missing, we simply use the value from the other year.

<sup>12</sup>In our data, consumption and total assets variables are never zero or negative, and so cause no problems for the log transformation. For other variables that may take zero values (such as subcategories of assets and investments), we add one before taking the log. Results are robust to alternative transformations of the dependent variable, such as the inverse hyperbolic sine transformation (IHST) of  $x$  is  $\log\left(x + (x^2 + 1)^{\frac{1}{2}}\right)$ , which unlike the log transformation is defined at zero (Burbidge et al. (1988).)

treatment singly. The complementarity parameter  $\gamma_b$  is statistically significantly different from zero at the 5% level. We reject the hypothesis that the two complementarity terms are jointly zero (p-value = 0.037).

Results from the assets regression (column 2) are broadly similar, but the coefficient estimates are less precise. Treatment effects on log assets are all positive, ranging from roughly 0.12 to 0.20. The treatment effects for the subsidy alone and basic savings alone are each statistically significantly different from zero at the 10% level. We cannot reject at conventional statistical significance levels that all treatment effects are equal in magnitude (p-value = 0.923), but also cannot reject that treatment effects are jointly zero (p-value = 0.519). As in the consumption regression, the complementarity parameters are both negative, and we find that basic savings and the subsidy are substitutes with respect to assets ( $\gamma_b$  is statistically significantly different from zero at the 10% level.)

Results in column 3 where the dependent variable is index of consumption and assets are consistent with the results for consumption or assets alone, with statistical significance levels more akin to the consumption results in column 1. Four out of five of the treatments have positive impacts that are statistically significantly different from zero at conventional levels. (The exception is the effect of the basic savings and subsidy treatment, which has a p-value of 0.125.) We do not reject at conventional levels that all treatment effects are equal in magnitude (p-value = 0.515), and do reject that all are jointly zero (p-value = 0.028). We also reject the hypothesis that the two complementarity terms are jointly zero with respect to this outcome variable (p-value = 0.016).

In sum: with respect to consumption, assets, or an index of these two outcomes, each of the five treatments, whether singly or jointly, has similar positive impacts. When subsidies are combined with either savings treatment, impacts are no larger than for the subsidy treatment by itself. Complementarity parameters are all negative in sign. This pattern leads us to characterize savings and subsidies as substitutes, rather than complements, with respect to the level of consumption and assets.

## 5 A theoretical model of the interaction between savings and agricultural technologies

The interventions offered to farm households in the study area can be characterized as follows:

- The *voucher coupon intervention* offered a 75% subsidy on the cost of an improved package of improved maize seeds and chemical fertilizers for the 2010/11 crop year. As we have shown in other work, voucher recipients substantially boosted their use of improved inputs, realized significant increases in maize production and increased their subjective beliefs about the returns to improved maize technology.
- The *simple savings intervention* made available mobile banking services that included access to a safe form of savings with a positive interest rate (x% per-annum) on savings held during the post-harvest to planting period and from the planting period to the next harvest. Training ...
- The *matched savings intervention* offered the same interest rate and security as the simple savings intervention for the planting to harvest period. In addition, it offered a substantially higher interest rate for savings during the post-harvest to planting season match period.

A three period household production/consumption model allows us to explore the core economic incentive effects created by the interventions whose impacts we study. In the initial post-harvest period, households must choose how much to consume and how much to carry forward for future consumption and agricultural

investment for the next season. In the planting season period, households must decide how much of the resources carried forward from the harvest season are consumed and how much are invested in next year's crop. Finally, in the terminal harvest period, households benefit from the net income that they have realized. Formally, we write this model as:

$$V_0(W_0, D_j^a, D_j^f | \alpha_0, \delta_0) \equiv \max_{c_t, S_t, K} u(c_0) + \beta u(c_1) + \beta^2 E_\theta [u(c_2)]$$

*subject to :*

$$\begin{aligned} c_0 &\leq W_0 - S_0 \\ c_1 &\leq (1 + \tilde{r}_1)S_0 - S_1 - pK \\ c_2 &\leq (1 + \tilde{r}_2)S_1 + (\bar{x} + \theta\tilde{\alpha}K) \\ S_0, S_1, K &\geq 0 \end{aligned}$$

where  $W_0$  is initial cash on hand post-harvest,  $r_1$  denotes the interest rate during the post-harvest period,  $r_2$  denotes the interest rate for the post-planting period and  $\tilde{\alpha}$  denotes subjective beliefs about the physical returns to improved agricultural inputs  $K$  which are purchased at price  $p$ . The price of the agricultural output has been normalized to one.

Absent the savings interventions, we assume that  $r_1 = r_2 = \delta \leq 0$ . This assumption of a non-positive return on savings can be thought of as reflecting deterioration of commodity savings, theft of insecurely held informal savings, and, or the inability of households to control either their own (Ashraf et al. (2006), Duflo et al. (2011), Dupas and Robinson (2013b), Gine et al. (2014)) or their neighbors demands for cash (Ashraf et al. (2015), Platteau (2000)). The simple savings intervention raises interest rates such that  $r_1 = r_2 = r_s > 0$ , where  $r_s$  is the standard bank savings rate. The matched savings intervention creates the interest rate structure  $r_1 = r_m > r_2 = r_s$ , where  $r_m$  is the interest rate created by the savings match.

We write the perceived returns to the agricultural technology as  $\bar{x} + \theta\tilde{\alpha}K$ , where  $\bar{x}$  is the returns to the traditional technology when no improved inputs are used,  $K$  is the amount invested in improved agricultural inputs. Returns to the improved technology are stochastic and the random variable  $\theta$  has base of support  $[\theta_{min}, \theta_{max}]$  and expected value equal to one. We assume that over the relevant range, returns to investment in the improved agricultural technology do not diminish.<sup>13</sup> Consistent with our data, we assume that absent further experimentation and learning, beliefs on the returns to the technology are downwardly biased such that  $\tilde{\alpha} = \alpha_0 + b$  where  $\alpha_0$  is the true returns to the technology and the bias  $b < 0$ .

This household problem is most easily solved by beginning with the second period first. Taking as given the amount of savings in the first period, we can write the planting season problem as a function of planting season cash on hand,  $W_1 = (1 + \tilde{r}_j)S_0$ :

$$V_1(W_1 | r_2, \tilde{\alpha}) \equiv \max_{c_t, S_t, K} u(c_1) + \beta E_\theta [u(c_2)]$$

---

<sup>13</sup>Explain this assumption ...

subject to :

$$\begin{aligned} c_1 &\leq (1+r_1)S_0 - S_1 - pK \\ c_2 &\leq (1+r_2)S_1 + (\bar{x} + \theta\tilde{\alpha}K) \\ S_1, K &\geq 0 \end{aligned}$$

The first order conditions with respect to  $S_1$  and  $K$  respectively are:

$$\begin{aligned} (1+r_2)\beta E(\theta u'_2) &\leq u'_1 \\ (\tilde{\alpha}/p)\beta E(\theta u'_2) &\leq u'_1 \end{aligned}$$

Note that  $u'_1$  on the right hand side of these inequalities essentially serves as the shadow cost of capital or liquidity. Pessimistic expectations about returns to the improved technology may make a corner solution with  $K = 0, S_1 > 0$  possible where discounted expected returns to investment do not exceed the cost of capital. Indeed, at the pre-intervention negative interest rate  $\delta$  impatience will surely hold (i.e.,  $(1+r_2)\beta < 1$ ) and the dual corner solution  $K, S_1 = 0$  could in turn easily hold for reasonable values of  $W_1$  and  $\bar{x}$ .

Inspection of the first order conditions make clear that a voucher subsidy which reduces  $p$  will make positive investment in  $K$  more likely. If that investment in turn reduces the bias in expected returns,  $\tilde{\alpha}$  will increase and may sustain investment in  $K$  even after the voucher subsidy ends and the input price  $p$  rises to its unsubsidized level. An interior solution for both choice variables, would be characterized by the following condition:

$$\frac{(\tilde{\alpha}/p)}{(1+r_2)} = \frac{E[u'_2]}{E[\theta u'_2]}.$$

Under the reasonable assumption that the true expected returns to investment exceed the rate of interest on formal savings ( $\alpha_0/p > (1+r_s)$ ), the hand side of this expression will be strictly greater than one. At the same time, assuming risk aversion, the right hand side of this equation will also be strictly greater than one for all positive values of  $K$  and will continue to further increase as  $K$  and the risk exposure of the household increase. Despite the gap in expected returns between the two investments,  $K$  and  $S_1$ , an interior solution is possible with both positive if the household chooses to diversify against the risk of investing in  $K$ . Note that the fraction  $1/(1+r_2)$  is the price of self-insurance through savings. When  $r_2 = 0$ , this insurance is actuarially fair (a dollar placed into savings returns a dollar), whereas values of  $r_2$  below (above) zero make the insurance actuarially unfair (favorable).

At this point, it is easy to see the impact of savings interventions that increase  $r_2$ . Such an increase first reduces the price of insurance through savings and will, other things equal, induce the household to buy more insurance and invest less. We denote this a substitution effect of a higher  $r_2$  as cheaper insurance leads to a substitution between riskier and safer investment.

On the hand, and again holding all else equal, the increase in  $r_2$  also reduces the correlation between  $\theta$  and  $u'_2$  and causes the right hand side of the expression to increase. This reduction in risk exposure will encourage the household to invest more in the productive, but risky investment  $K$ . We call this the risk-bearing effect of a higher  $r_2$ . In general, there is no way to sign whether or not the net effect of an increase in  $r_2$  will bring an increase or a decrease in investment in  $K$ . However, under a wide range of assumptions, the substitution effect will dominate.<sup>14</sup> We will return to this issue in an illustrative numerical analysis below.

<sup>14</sup>Intuitively, the substitution effect will tend to dominate because households will tend to be woefully underinsured when  $r_2$

Importantly, note that this substitution into savings/insurance will reduce realized consumption variability.

Finally note that the second period value function is non-decreasing in  $r_2$  and  $\tilde{\alpha}$  (increasing for interior solutions).

Using the value function  $V_1(W_1)$  defined by the second period problem, we can now rewrite the full three period problem as:

$$V_0(W_0) \equiv \max_{c_0, S_0} u(c_0) + \beta V_1(W_1 | r_2, \tilde{\alpha})$$

*subject to :*

$$c_0 \leq W_0 - S_0$$

$$W_1 = (1 + r_1)S_0$$

$$S_0 \geq 0$$

This problem implies the following first order condition:

$$u'_0 \geq (1 + r_1)\beta \frac{\partial V_1}{\partial W_1}.$$

As this condition makes clear, an increase in the post-harvest interest rate will (assuming an interior solution with  $S_0 > 0$ ) increase planting season cash on hand  $W_1$ . Holding other things equal, this increase in  $W_1$  will lower the shadow price of liquidity ( $u'_1$ ) and potentially boost investment in both  $S$  and  $K$  via this wealth effect. By cheapening the cost of investment, an increase in the post-harvest interest rate  $r_1$  operates exactly like an input subsidy. In the case of the matched saving intervention (where the 4-month post-harvest interest rate rose to 25%), this subsidy-equivalent effect was quite substantial, although still less than the 75% input price reduction offered by the voucher program.

\*\*\*\*

We are now in a position to consider the impacts of the different interventions. Given the number of moving parts, this is most easily seen through numerical analysis of the model developed here. Consistent with what the program was designed to do, we make the following specific assumptions about the parameters for the different treatment groups in the years following the voucher coupon intervention.

To bring out the impacts that the model above would result from these treatments, we focus on two outcomes measures: the expected value of terminal period consumption,  $c_3$ ; and, the variance of terminal period consumption.

\*\*\*\*

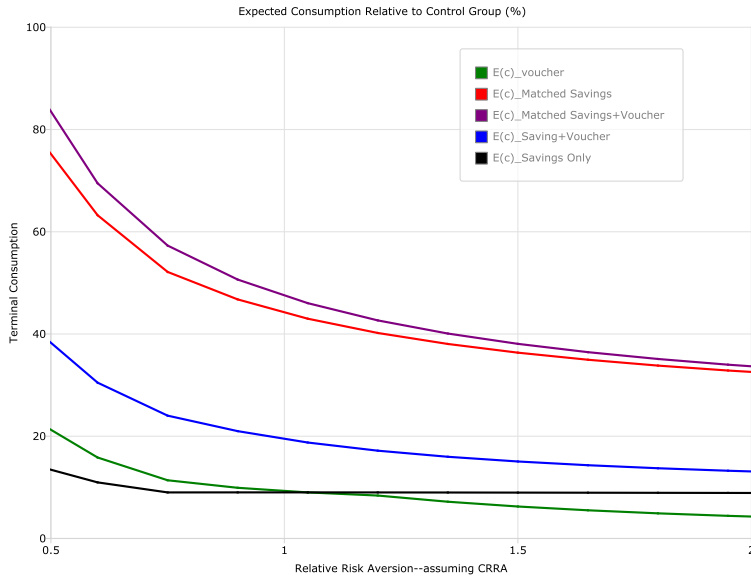
The top panel in Figure 1 compares the model's predictions about the impact of the voucher only treatment with the impacts of the different matched savings treatment. The horizontal axis displays different levels of risk aversion, while the vertical axis displays key impact indicators as a percentage of the indicator's value for the control group. For each treatment group, two curves are graphed: expected terminal consumption (solid lines) as well as the coefficient of variation of terminal consumption (dashed lines). <<DEAN AND RACHID, WE COULD ALSO GRAPH TOTAL VARIANCE IF YOU PREFER>>. Comparing just voucher only treatment with the matched savings only treatment we see that matched savings amplifies somewhat

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is low .... .

Figure 1: Results of Numerical Analysis

(a) Impacts on Consumption Level



(b) Impacts on Consumption Variability

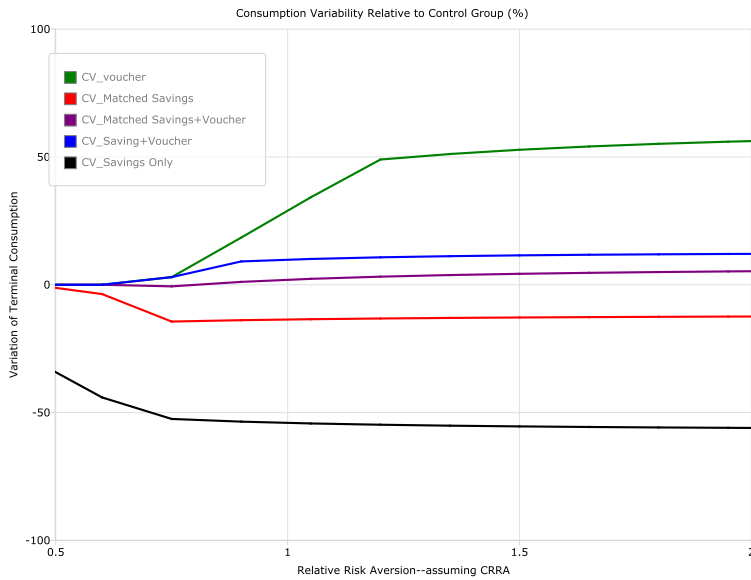


Table 1: Parameter Values Used for the Numerical Analysis

<i>Treatment Groups</i>	<i>Parameters</i>		
	Expected Returns, $\bar{\alpha}/p$	Post-harvest Interest Rate, $r_1$	Post-planting Interest Rate, $r_2$
Control	110%	-4%	-4%
Voucher Only	125%	-4%	-4%
Matched Savings Only	120%	25%	4%
Matched Savings + Voucher	125%	25%	4%
Simple Savings Only	110%	4%	4%
Simple Savings + Voucher	125%	4%	4%

Constant relative risk aversion preferences and a per-period discount factor of 0.95  
List numerical distributional assumption on theta

the investment and expected consumption impacts of the voucher only treatment (recall that these results assume that the matched savings encourages experimentation and results and increases subjective returns almost as much as the voucher treatment alone. In addition, the increase in the post-planting interest rate allows households to cheaply insure themselves against the increased consumption variability. Relative to the control group, the coefficient of variation of consumption drops for the matched savings group, whereas it rises sharply the voucher only group. We thus see that while matched savings (at least when combined with training that primed farmers to invest in improved agricultural inputs) is in some ways a substitute for the voucher coupons, it is in other way a complement as it allows expected consumption increases without increased, utility-diminishing consumption variability.

Turning to the bottom panel in Figure 1, we see that when combined with the voucher coupon, simple savings operates much like matched savings, although income increases are less sharp as it remains relatively expensive to move money from the harvest to the planting period. Simple savings by itself has a more modest impact on expected consumption than do vouchers alone when risk aversion is low. However, as risk aversion increases above about 2, simple savings alone results in higher expected consumption than vouchers alone. For these highly risk averse households, the costliness of self-insurance with improved savings mechanisms discourages investment even in a technology understood to be profitable.

## 6 Additional analyses

In this section we conduct additional analyses to test auxiliary predictions of the theory.

### 6.1 Impacts on the variance of consumption

In theory, subsidies and savings may be substitutes (from the standpoint of consumption) if households are optimizing by jointly choosing the level and variance of consumption. Households could respond to a single treatment by taking actions (such as increased fertilizer use) to increase the level of consumption. When offered a combination of treatments, they might not seek to further increase consumption levels (beyond what a single treatment allows), and instead seek to reduce consumption variance. This is possible because savings can play a dual role in household financial decision-making: savings can facilitate increasing consumption levels (by facilitating accumulation of resources for investment) and also help reduce variance (by facilitating buffer stock accumulation and possibly other simultaneous changes in intertemporal activities).



Evidence in support of this explanation would be that the subsidy treatment increases consumption variance, but by less when combined with the savings treatments. This turns out to be the case.

We provide formal tests of differences in variance across the treatment groups in Table 4. In column 1, the table displays, in the top six rows, the standard deviation of log consumption (the dependent variable in column 1 of Table 3) in the control group and each of the separate treatment groups. In the bottom part of the table, we report p-values of F-tests of equality of variances across pairs of treatment groups.

The subsidy treatment, by itself, leads to higher consumption variance ( $\sigma_v = 0.545$ ) compared to the pure control group ( $\sigma_c = 0.453$ ), and this difference is statistically significant at conventional levels (p-value = 0.004, top row of Panel A). Also, consumption variance in the basic savings only group is statistically significantly higher than in the control group (p-value = 0.093). By contrast, consumption variances in the remaining treatment groups are all lower than in the subsidy-only treatment group, and none are statistically significantly different from the variance in the control group.

When combined with either savings treatment, the subsidy leads to less consumption variance than the subsidy treatment by itself. P-values in Panel B are for pairwise variance comparisons of the subsidy treatment alone versus other treatments that include savings. Rows 2 and 4 of the panel indicate that, compared to the subsidy-only treatment, consumption variance is lower in the basic savings and subsidy treatment ( $\sigma_{bv} = 0.480$ , p-value = 0.039) and in the matched savings and subsidy treatment ( $\sigma_{mv} = 0.446$ , p-value = 0.002).

It also appears that the matched savings treatment, by itself, leads to lower consumption variance ( $\sigma_m = 0.489$ ) than the subsidy treatment by itself (p-value = 0.098).

Across all the treatment groups, the matched savings and subsidy treatment leads to the lowest consumption variance. Consumption variance in that treatment is statistically significantly lower than consumption variance in the basic savings only treatment (p-value = 0.064) and the subsidy-only treatment (as mentioned above), but the difference is not statistically significant at conventional levels vis-a-vis any of the other treatments.

The reduction in the variance of consumption when the subsidy treatment is combined with the savings treatments (compared to the subsidy-only treatment) can also be seen graphically in Figure 1. The figure displays probability density functions of log consumption (the same dependent variable as in column 1 of Table 3) for the three subsidy treatment groups. PDFs for the subsidy treatments combined with either the basic savings (long dashed line) or matched savings (short dashed line) groups are less dispersed about their respective means, compared with the PDF of the subsidy-only group (solid line).

In sum, the subsidy treatment, when offered by itself, leads to a statistically significant increase in consumption variance, compared to the control group. By contrast, consumption variances in the joint savings and subsidy treatments are statistically significantly lower than in the subsidy-only treatment, and are statistically indistinguishable from the level of consumption variance in the control group.<sup>15</sup> The savings treatments appear to help subsidy recipients offset the additional risk associated with increased agricultural input investments.

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<sup>15</sup>One can also compare the conditional variance of log consumption across treatment conditions, after conditioning on the full set of control variables in equation 1. The conclusions of this section related to differences in unconditional variance are highly robust to examining conditional variances instead (which is unsurprising, given randomization of treatments and balance on key baseline covariates).

## 6.2 Impacts on quantiles of the consumption distribution

A complementary view of the impacts of the treatment effects is provided by examining impacts on quantiles of the consumption distribution. Table 5 presents quantile regressions for deciles of the consumption distribution, from the 10th percentile (column 1) to the 90th percentile (column 9). The regression's right-hand-side variables are identical to those of equation 1. Each coefficient presented represents the effect of the given treatment combination on the percentile of the log per capita consumption distribution. Towards the bottom of the table, we report per capita consumption (in Mozambican meticals) at the given percentile in the control group (which ranges from MZN 36 at the 10th percentile, to MZN 62 at the median, and MZN 114 at the 90th percentile).

The broad pattern in these results is that all treatments clearly have large positive impacts at the 90th percentile of the distribution. A key pattern of difference across the treatments is that the subsidy-only treatment shows little effect at lower percentiles, while the savings treatments tend to show some positive impacts at lower percentiles.

The subsidy-only treatment quite clearly has its primary impact at the upper deciles of the distribution, with an effect at the 90th percentile of roughly 18 log points (which is statistically significant at the 1% level). Point estimates at the 60th to 80th percentiles are nontrivial in magnitude (ranging from 6 to 9 log points), but are not statistically significantly different from zero.

The other treatments that involve savings also have substantial positive impacts on the 90th percentile, with point estimates ranging from 11 to 18 log points, and all but one (for basic savings plus subsidy) are significant at the 5% level or better. Broadly, the savings treatments are more likely to show evidence of effects at lower percentiles. The matched savings plus subsidy treatment has positive coefficients that are nontrivial in magnitude for all percentiles below the 90th (ranging from 6 to 12 log points), and those on the 20th, 40th, 50th, and 70th are statistically significant at conventional levels. Coefficients for matched savings only are positive and roughly 7 log points for all percentiles from the median and below (although only statistically significant at the 20th percentile). Coefficients for basic savings only are 7 log points or larger for the 30th through 80th percentiles, and statistically significantly different from zero for the 40th through 60th percentiles. Point estimates for the basic savings plus subsidy treatment are smaller, but still nontrivial at some lower percentiles: coefficients for the 20th to 50th percentiles range from 4-6 log points, although none are statistically significant at conventional levels.

So far, we have identified two broad patterns in the results. First, the savings treatments lead to lower consumption variance than does the subsidy-only treatment. Second, positive consumption impacts are more likely to extend to lower percentiles of the distribution for the savings treatments than for the subsidy-only treatment. These patterns are consistent with two explanations. On the one hand, they are consistent with improved risk management on the part of study participants in the savings treatments: buffer stocks of savings could be helping households keep their consumption from falling too low in the face of negative shocks. On the other hand, however, they could mean that the savings treatments do a better job of targeting poorer households in the sample, in comparison to the subsidy-only treatment that differentially benefits the upper tail of the distribution. We seek evidence mediating between these two explanations in the next two subsections.

### 6.3 Treatment effect heterogeneity with respect to baseline consumption

An open question is whether the patterns identified in subsections 6.1 and 6.2 above (wherein the savings treatments lead to lower consumption variance and to treatment effects on consumption that extend lower in the consumption distribution) reflect improvements in risk management or, alternately, that the savings treatments are more “pro-poor” (provide differentially more benefit to initially poorer households, compared to the subsidy-only treatment). To examine whether the savings treatments appear to be more pro-poor, we test for heterogeneity in treatment effects with respect to baseline consumption (log per capita consumption in the household April 2011).<sup>16</sup> Baseline consumption is normalized to have zero mean and unit standard deviation, to ease interpretation of the main effects of treatment.

Results are in Table 6. All regressions contain the same control variables and fixed effects as equation 1. In column 1, the right-hand-side variables of interest are simply an indicator for being in any treatment group, and its interaction with baseline consumption. The main effect of baseline consumption is also included in the regression.<sup>17</sup> Because the baseline consumption variable has been demeaned, the coefficient on “Any treatment” is the treatment effect for a household at mean baseline consumption, estimated at 7.0 log points, which is significant at the 1% level. The interaction term is positive and significant at the 10% level, indicating that the treatment effect is larger for households with higher baseline consumption. A one-standard-deviation increase in baseline consumption leads the treatment effect to be larger by 4.2 log points. The full set of treatments, on average, are not in fact pro-poor: households with higher consumption at baseline benefit more from the treatments.

We now test whether the subsidy-only treatment is more or less pro-poor than the savings treatments, on average. The independent variables of interest in the next column in the table are indicators for the subsidy only treatment, indicators for having received any of the savings treatments, and the interactions of these treatment indicators with baseline consumption. The coefficients on the main effects of the treatments each indicate positive and statistically significant effects for households with mean baseline consumption. The coefficients on the interaction terms with baseline consumption are remarkably similar to one another: 0.045 and 0.042 for the interactions with, respectively, the subsidy-only indicator and the “Any savings treatment” indicator. Only the latter coefficient is statistically significantly different from zero (at the 10% level). The key test, however, is whether these two coefficients are statistically significantly different from one another. They are not: an F-test (reported at the bottom of the table) does not reject equality of the coefficients (p-value = 0.952). (An F-test also does not reject the hypothesis that both interaction terms are zero.)

Column 3 of the table is similar to column 2, but it now includes main effects and interactions with the complete set of treatment combinations. The results are similar: F-tests do not reject equality of the coefficients on the interaction terms with baseline consumption, and also do not reject that all these interaction term coefficients are zero.

In sum, we find no indication that the savings treatments are more pro-poor than the subsidy-only treatment. This analysis suggests that the patterns identified in subsections 6.1 and 6.2 above are not likely to be due to the savings treatments providing differentially greater benefit to initially poorer households,

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<sup>16</sup>The April 2011 survey is a baseline with respect to the savings treatments, which were offered just after that survey was collected. That survey is however not a true baseline with respect to the subsidy treatment, since the subsidies were provided in November 2010. That said, April 2011 is just before the 2011 harvest season (May-July), so household consumption should not have been directly affected yet by the subsidies. We do confirm that log per capita household consumption in the April 2011 survey is uncorrelated with the subsidy only treatment or any of the other treatment combinations (results available upon request.)

<sup>17</sup>As would be expected, in each of the regressions baseline consumption is positively and statistically significantly associated with the dependent variable.

and so indirectly support the risk-management interpretation of those results. The next subsection provides more direct support of the risk-management interpretation.

## 6.4 Consumption smoothing in the face of shocks

More direct evidence in favor of the risk-management interpretation of the patterns in subsections 6.1 and 6.2 would be if households in the savings treatments were better able to insulate consumption from the effect of negative income shocks. Regressions in Table 7 investigate this question. These regressions are identical in structure to those of Table 6 in the preceding subsection, but where we replace the “Baseline consumption” variable and interactions with a variable measuring agricultural shocks, “Bad years”. This variable counts how many of the agricultural seasons of 2012 and 2013 (in which harvests were roughly in May to July of those years) were “very bad” for agriculture, self-reported by the respondent in those years’ follow-up surveys (September 2012 and August 2013).<sup>18</sup> It takes on values of 0 (59.4% of respondents), 1 (34.46%), or 2 (6.19%).<sup>19</sup>

The main effect of “Bad years” is interpreted as the impact of one additional bad year in 2012-13 on mean log consumption in 2012-13. The coefficient on the interaction between “Bad years” and a treatment indicator represents the impact of the specified treatment on the impact of a bad year on consumption. A negative coefficient would mean that a treatment makes a bad year even worse for consumption (it magnifies risk), while a positive interaction term coefficient would mean the opposite: the treatment attenuates the impact of a bad year on consumption (a reduction in risk). The key test is whether the interaction term on the subsidy-only treatment is more negative than the interaction term on the savings treatments (whether the latter minus the former is positive). If so, this would be consistent with the savings treatments providing better risk management, compared to the subsidy-only treatment.

This does seem to be the case in the regressions of Table 7. The first column, in which the treatment variable is simply an indicator for “Any treatment”, is shown as a point of comparison. Any treatment leads to higher consumption. The coefficient on “Any treatment” interacted with bad years is actually positive, but is small in magnitude and is not statistically significantly different from zero. The test of the hypothesis comes in column 2, which includes as independent variables the subsidy-only treatment indicator, an indicator for “Any savings treatment”, and their interactions with bad years. The coefficient on the interaction with the subsidy-only treatment is negative, while the coefficient on the interaction with any savings treatment is positive. While neither interaction term coefficient is statistically significant on its own, the test is whether the two are different from one another. Indeed, they are, as shown in the F-test reported at the bottom of the table, which rejects (p-value = 0.007) that the two coefficients are equal to one another. An F-test also rejects the hypothesis that the two are jointly zero (p-value = 0.025).

The results in column 3 provide further detail by interacting each treatment combination with bad years, and come to the same conclusion. The interaction term with subsidy-only is negative, and each coefficient on an interaction term with a savings treatment is positive. We reject at conventional levels (p-value = 0.067) the hypothesis that all interaction term coefficients are equal.<sup>20</sup>

<sup>18</sup>After a set of questions asking respondents to estimate the returns to fertilizer in an “average year”, a “very good year”, and a “very bad year”, the respondent is asked “How would you consider the current year?” Possible responses are “very good”, “very bad”, and “regular”.

<sup>19</sup>One might worry that the occurrence of bad years might be endogenous to treatment. However, we find that respondent treatment status is not correlated with their “Bad years” variable. When regressing “Bad years” on indicators for each treatment combination (and all control variables used in the main results tables), no coefficient on a treatment indicator is large or statistically significantly different from zero, and an F-test does not reject that the treatment coefficients are jointly zero (p-value 0.478).

<sup>20</sup>Results are robust to alternative specifications of the results in these “Bad years” regressions, such as running regressions

In sum, consumption appears to be better insulated from the negative effects of bad agricultural shocks in the savings treatment groups than in the subsidy-only treatment group. This is consistent with better risk-management on the part of respondents receiving the savings treatments, and so helps support the risk-management interpretation of the reduction in consumption variance in the savings treatment groups (compared to the subsidy-only group) and the fact that the positive impacts on consumption per capita extend to lower percentiles of the consumption distribution in the savings treatment groups than in the subsidy-only treatment group (results in subsections 6.1 and 6.2 respectively).

## 6.5 Impacts on investment activities

We now turn to understanding what intermediate outcomes may have mediated the impacts on consumption per capita and on risk management. Changes in household investments may have mediated impacts on consumption, and adjustments to asset holdings may also play a role in risk management. We therefore examine treatment effects on a variety of different types of investment activities. Results from estimation of equation 1 are presented in Table 8.

Given the emphasis on promoting use of modern agricultural inputs in all the subsidy and savings treatments, we first examine impacts on log fertilizer purchases in column 1. With the exception of the basic savings-only treatment, the treatment coefficients are positive and large in magnitude. Effects of the subsidy only, basic savings plus subsidy, and matched savings only treatments are statistically significantly different from zero at conventional levels. (The matched savings plus subsidy treatment comes close to statistical significance, with a p-value of 0.115.) For the basic savings-only treatment, on the other hand, the point estimate is much smaller in magnitude and is not statistically significantly different from zero. This pattern is unsurprising, given that the basic savings-only treatment was the only treatment that did not provide a generous financial incentive for fertilizer use. An F-test does reject (with p-value 0.018) that all treatment effects on fertilizer are jointly zero. It appears that all treatments that explicitly incentivized fertilizer use were successful in doing so.

There is modest evidence of impacts on other types of investment. Treatment effects on education, most types of agricultural investment, and on non-agricultural investments are uniformly not statistically significantly different from zero. The exception to this is impacts on log “other” agricultural investments (column 6): the basic savings-only treatment and the combined matched savings and subsidy treatment both have positive effects on this outcome that are statistically significant at the 5% level.<sup>21</sup>

As an addendum, we also show treatment effects on log crop production (column 9). While we reject at conventional significance levels (p-value = 0.062) that all treatment effects are zero, there is heterogeneity in the treatment effects (the F-test rejects that all treatment effects are equal, with a p-value of exactly 0.060). While all treatment effects are positive, only the subsidy-only treatment and the matched savings-only treatment have economically and statistically significant impacts (at the 10% and 5% levels, respectively). For the other treatments, coefficient estimates are much smaller in magnitude and are not statistically significantly different from zero. This heterogeneity in impacts on crop production may in part reflect the riskiness of crop production, with some areas experiencing locality-specific negative shocks. That we nonetheless estimate relatively uniform impacts across the treatments on consumption and assets may reflect

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solely using 2013 values of consumption and bad year, solely using 2012 values of consumption and bad year, and pooling annual observations for 2012 and 2013 into one regression (instead of taking averages across 2012-13).

<sup>21</sup>That said, we note that these effects are small in absolute value, since they are impacts on log investment with respect a very small average in the control group of just 42 MZN.

that households are generating income from sources other than crop production that we do not observe.<sup>22</sup>

## 6.6 Impacts on financial management (asset allocations and loans)

The savings treatments were aimed at encouraging households to save in formal banks. This may have led to changes in asset allocations, from non-financial or non-bank savings towards savings in formal institutions. Increased savings in formal banks may have also affected access to credit. We refer to management of such household balance sheet items (borrowing and specific subcategories of assets) as “financial management” activities. (Note that livestock may also serve an investment or production role.) Changes in these activities may be channels through which the treatments led to increases in consumption, and may also play a role in risk management.

Impacts on financial management activities, estimated using equation 1, are presented in Table 9. The most prominent treatment effects in the table are on formal savings (column 1). Each treatment has large, positive, and statistically significant effects on formal savings balances. An F-test rejects strongly that all treatment effects are zero (p-value = 0.000.) There is heterogeneity in the treatment effect: an F-test also rejects the null that all treatment effects are equal (p-value = 0.049.) This appears to be driven by the impact of the subsidy-only treatment (itself statistically significant, at the 5% level) being much smaller in magnitude than the other treatment effects. The impacts of the savings-only and savings plus subsidy treatments are all larger in magnitude and statistically significantly different from zero at the 1% level.<sup>23</sup> This heterogeneity in impacts on formal savings is unsurprising, given that the subsidy-only treatment did not involve any facilitation of formal savings.

Evidence of impacts is more tenuous for the other asset categories (in columns 2-5), perhaps with the exception of cash held outside of banks (column 2). For this outcome, only one of the five treatment effects is statistically significantly different from zero at conventional levels (the coefficient on the matched savings only treatment, at the 5% significance level), but four out of five are positive and large in magnitude (with the exception of the basic savings plus subsidy treatment), and an F-test rejects at conventional levels that the treatment effects are jointly zero (p-value = 0.080). For this outcome we also reject that the complementarity parameters are jointly zero (p-value = 0.094).

Impacts on loan balances are displayed in column 6. All treatment effect estimates are positive, and the effect is statistically significantly different from zero for the basic savings and subsidy treatment (at the 5% level). However, we cannot reject that all treatment effects are zero (p-value = 0.170.)<sup>24</sup>

## 6.7 Impacts on indices of intertemporal activities

In Tables 8 and 9, we examined impacts on several dependent variables, which gives rise to concerns related to multiple inference. In Table 10, therefore, we estimate equation 1 where the dependent variables are

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<sup>22</sup>While we examine non-agricultural investments in columns 7 and 8, and find little evidence of impacts, our surveys are relatively sparse on non-agricultural activities (and have no questions on non-agricultural income).

<sup>23</sup>Pairwise F-tests of the equality of treatment effects (not reported in the table) reveal that the effect of the subsidy-only treatment is different from each of the other treatment effects at conventional significance levels (at the 5% level for each pairwise comparison with the basic savings-only, basic savings plus subsidy, and matched savings-only treatments, and at the 1% level in the pairwise comparison with the matched savings plus subsidy treatment.)

<sup>24</sup>Morduch (2009) discusses theories that might explain simultaneously borrowing and saving (which he and co-authors also observe with frequency in the Collins et al. (2009) financial diaries.) Individuals holding savings stocks (which may be for buffer stock or asset accumulation purposes) may borrow to avoid drawing down these stocks. Due to high interest rates and enforcement from formal lenders, there is high pressure to repay credit relatively quickly, but less pressure to build savings back up if depleted. Individuals anticipating that self-control problems will delay their re-accumulation of savings may choose to take our loans instead of drawing down savings when an immediate liquidity need arises.

indices of sets of outcomes (Kling et al. (2007).) The dependent variable in column 1 is an index of the eight investment activities examined in Table 8 (not including crop production, which is an outcome of investment rather than an investment itself.) In column 2, we examine the index of the six financial management outcomes in Table 9, and in column 3 the dependent variable is an index of all the outcomes in Tables 8 and 9 (again excluding crop production), which we refer to as an “index of intertemporal activities.” In each case, the index is the average of all the relevant variables, after each variable has been normalized by the mean and standard deviation in the control group.

Results in column 1 provide no more than weak indication of impacts on investment activities overall. Only one out of five treatment effects is statistically significantly different from zero (the coefficient on the subsidy-only treatment, at the 10% level.) The other treatment effects are also positive but none are statistically significant at conventional levels. We cannot reject that all treatment effects are jointly zero (p-value = 0.571).

By contrast, there is strong evidence of impacts on financial management activities, overall. In column 2, each of five treatment effects is positive, large in magnitude, and statistically significant at the 1% level. An F-test rejects the null that all treatment effects are jointly zero (p-value < 0.001). This finding also holds, broadly, when examining the index of all intertemporal activities in column 3: each of the five treatment effects is statistically significantly different from zero at conventional levels (either the 5% or 1% level), and we reject the null that all treatment effects are jointly zero (p-value = 0.015).

Across columns 1 to 3, there is no indication that the treatment effects are different from one another: none of the F-tests can reject the null that all treatment effects are equal at conventional levels of statistical significance. In columns 2 and 3, we reject at conventional levels the null that all treatment effects are zero, as well as the null that the complementarity parameters (which are all negative in sign) are jointly zero.

In sum, there is strong evidence that the treatments had significant impacts on a broad range of intertemporal activities in study households. These overall effects are driven by changes in the set of financial management activities (credit and subtypes of asset holdings), rather than changes in the set of investment outcomes.

## 6.8 No evidence of impacts on transfers to other households

An open question is whether any of the effects on risk management are due to changes in informal insurance arrangements, in which households make transfers to one another to help smooth consumption. Two questions in the follow-up surveys help reveal whether the treatments change the extent to which study participants share resources with other households. The first question asks, “In the last three months, how many times have you been asked for money/help from someone who is not from your household?”, and is followed by “Out of these times, how many times did you help?” From answers in the 2012 and 2013 surveys, we construct two dependent variables: 1) an indicator for the respondent reporting to have assisted another household in either of those surveys, and 2) the total number of times the respondent reported assisted another household in those surveys (summed across the two survey rounds).

In Appendix Table 2, we report results from regressing these two dependent variables on indicator variables for each of the five treatment conditions. If changes in transfers were one mechanism through which the changes in consumption variance occurred, we would expect a positive coefficient on the subsidy-only indicator (increases in transfers to other households), and negative coefficients on the indicators for the savings treatments (decreases in transfers to other households). As it turns out, none of the coefficients are statistically significantly different from zero, and we also do not reject that they are jointly statistically

significantly different from zero. These results provide no indication that changes in informal insurance are in part responsible for the observed changes in consumption variance across treatments.

## 7 Conclusion

We conducted a randomized controlled trial to shed light on potential complementarities between modern input subsidies and savings facilitation programs. From the standpoint of raising consumption, subsidies and savings appear to be substitutes, rather than complements. Either treatment on its own has similar positive impacts, but providing both treatments has no larger impact. Savings treatments, when accompanying subsidies, instead facilitate risk-reduction, after households have taken (potentially risky) steps to raise consumption. Households receiving the savings treatments have lower consumption variance, and their consumption falls by less when they experience negative agricultural shocks. The results underscore how households seek to balance risk and return in their intertemporal decision-making, and reveal that complementarities between financial services interventions and other development programs may show up in reducing risk rather than in raising returns.

In the context of shedding light on the interplay among components of bundled development programs, we highlight the empirical relevance of the case of high substitutability between programs, so that implementation of one inexpensive program may have just as much impact (key summary measures of well-being, such as consumption) as combining that program with a more expensive one. In particular, our results emphasize the particular value of basic financial services such as savings for the risk-management benefits they bring. Households that have used subsidies to take (potentially risky) steps to raise consumption and assets may have limited appetite for further increasing their input use (and their concomitant risk exposure), and instead at that point may prioritize risk-reduction. In other words, for risk averse decision-makers, subsidy and savings interventions may be complementary (or at least additive) from the standpoint of raising expected utility, rather than the expected value of consumption.

An implication of our results is that input subsidy programs may be observed to have quite different impacts on input use, across environments that vary in their financial market development and the presence of other development programs. From the standpoint of raising fertilizer utilization, for example, we found that random assignment to the subsidy had large impacts in the no-savings-program localities, but no incremental impact on fertilizer use in localities that received the matched savings program. In matched savings localities, subsidy voucher winners and losers saw similar increases in fertilizer use (relative to the pure control group in no-savings-program villages). In the context of our overall results, this likely reflects the fact that the matched savings program provided the resources that allowed voucher losers to also raise their fertilizer usage, while voucher winners in the matched savings localities did not seek to further expand fertilizer use with the matched savings resources, perhaps due to the increased risk exposure this entailed. In other words, subsidies might be found to have little or no effect on fertilizer use if households are in relatively liquidity-unconstrained environments and have undertaken risky investments to such a degree that they have limited appetite for taking on more risk.<sup>25</sup>

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<sup>25</sup>This insight may help explain differences in the observed persistence of impacts of subsidies on fertilizer use across different studies. For example, Duflo et al. (2011) find no persistence beyond the subsidized season, in contrast to our findings of persistence (elaborated further in our companion paper, Carter et al. (2014).) It may be that households in the Kenyan environment of Duflo et al. (2011)'s sample are less liquidity-constrained and have reached a point where they would not want to take on more risk for higher returns (perhaps akin to the households in our matched savings villages).



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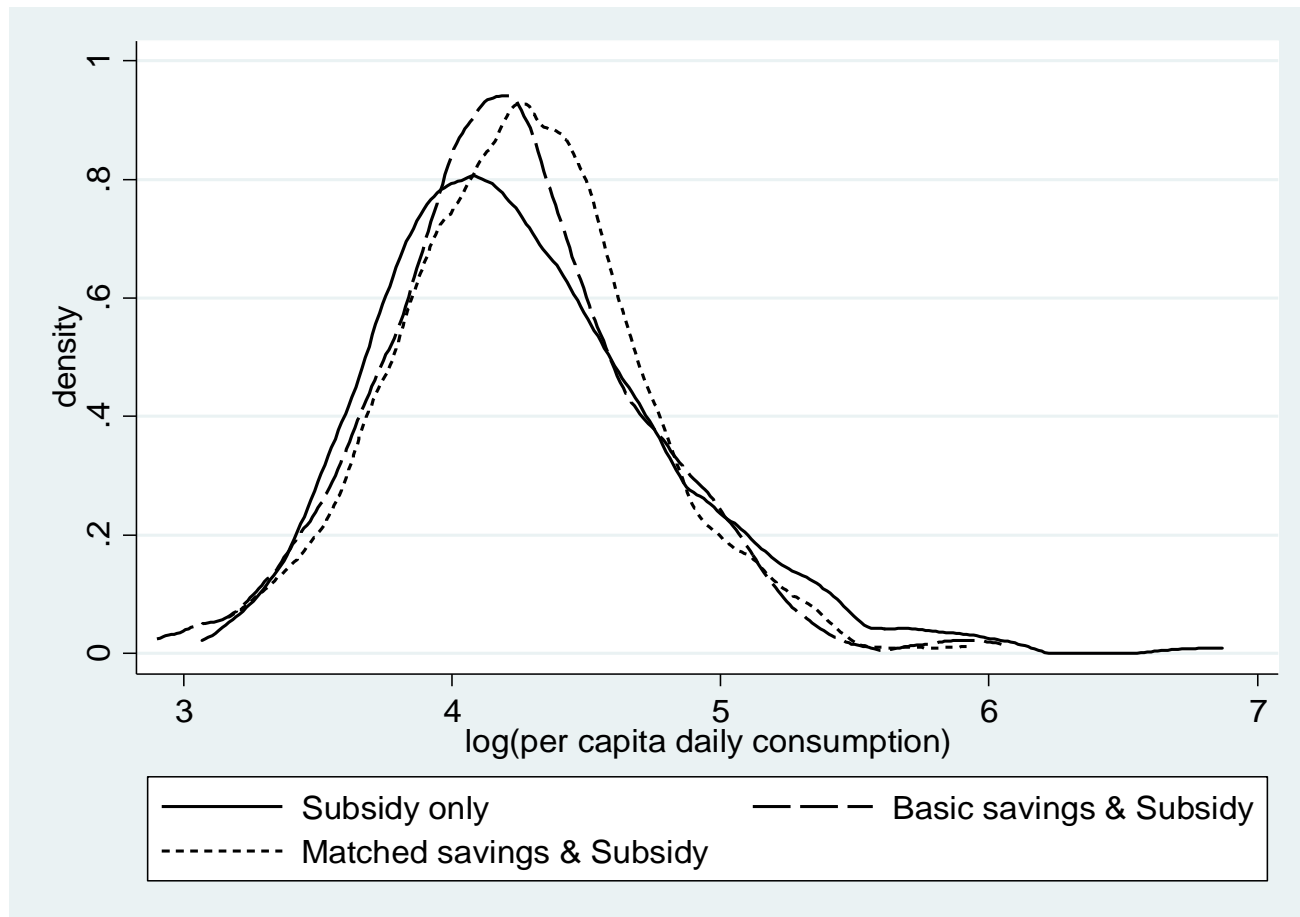
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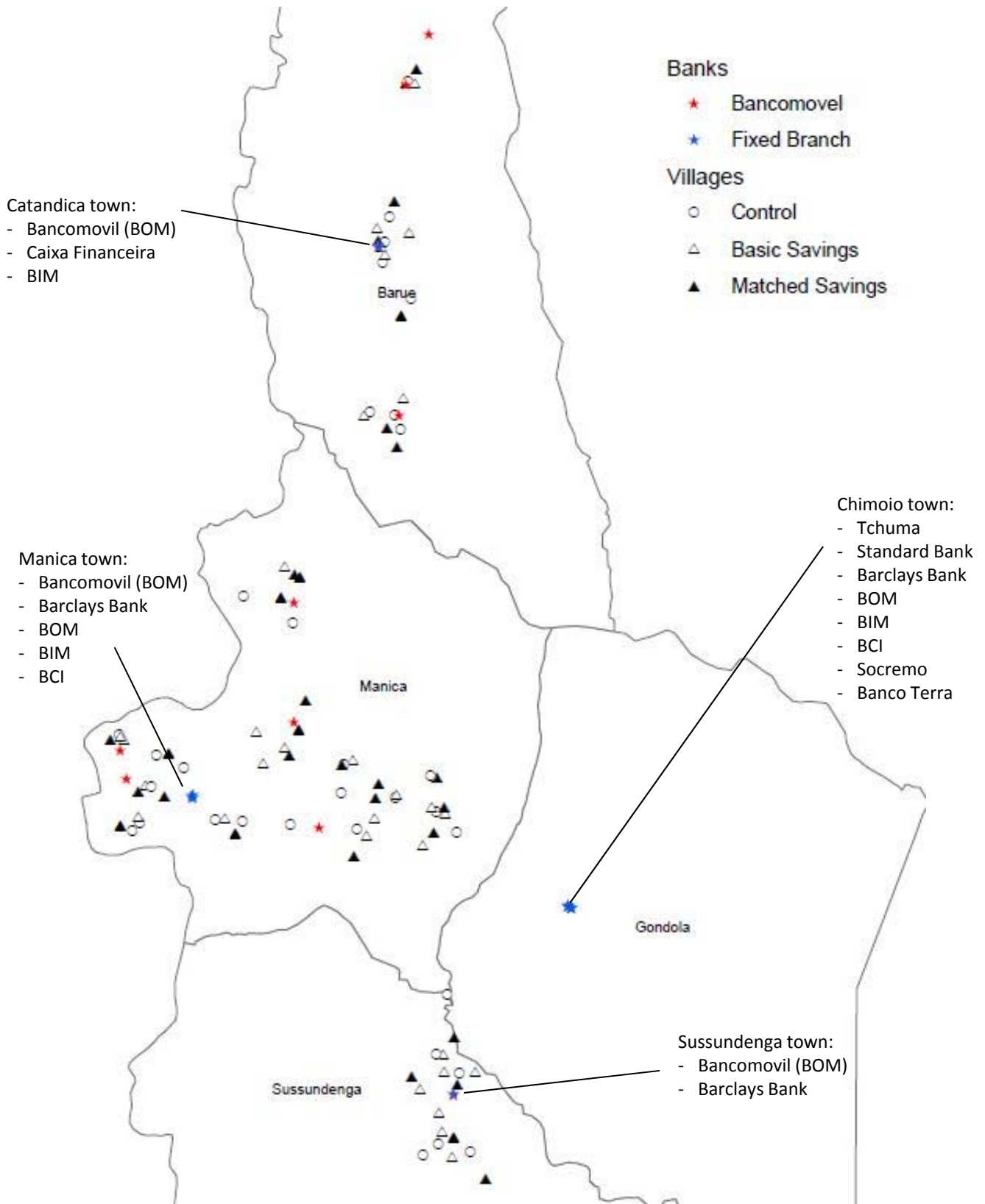
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Figure 1: Probability density functions of post-treatment log (consumption), subsidy treatment groups



Notes: Densities use Epanechnikov kernel with bandwidth 0.1. Per capita daily consumption is average value across Sep 2012 and Jul-Aug 2013 surveys. Treatments defined in Table 1 (T1: subsidy only; T3: basic savings & subsidy; T5: matched savings & subsidy). Subsidies randomly assigned in late 2010. Savings programs randomly assigned in April 2011.

**Figure 2: Study localities by savings treatment status, with bank locations**



Note: Borders demarcate districts within Manica province.

**Table 1: Treatment conditions**

		<u>Voucher treatment</u>	
		<i>No</i>	<i>Yes</i>
<u>Savings treatments</u>	<i>Control</i>	<b>C:</b> Pure Control (N=258)	<b>T1:</b> Subsidy (N=238)
	<i>Information</i>	<b>T2:</b> Basic savings (N=269)	<b>T3:</b> Subsidy & Basic savings (N=296)
	<i>Match</i>	<b>T4:</b> Matched savings (N=236)	<b>T5:</b> Subsidy & Matched savings (N=237)

Notes: Subsidy vouchers for agricultural inputs distributed one time, at start of 2010-2011 agricultural season (Sep-Dec 2010). Savings treatments administered in Mar-Jul 2011. Matched savings treatment provides temporary high interest rates in Aug-Oct 2011 and Aug-Oct 2012. Savings treatment conditions randomized across 94 study localities, each with 1/3 probability (32 control, 30 basic savings, 32 matched savings localities). Subsidy vouchers randomized at individual level (with 50% probability) within each study locality. Number of individual observations in parentheses. Total N=1,534.



**Table 2: Summary Statistics and Balance Tests**

	<i>Section A: Variables in levels</i>						<i>Section B: Continuous production variables in logs</i>					
	C: Pure Control	T1: Voucher	T2: Basic savings	T3: Voucher & Basic savings	T4: Matched savings	T5: Voucher & Matched savings	C: Pure Control	T1: Voucher	T2: Basic savings	T3: Voucher & Basic savings	T4: Matched savings	T5: Voucher & Matched savings
HH head education (yrs.)	4.77 (3.32)	4.7 (3.01)	4.75 (3.41)	4.83 (3.42)	4.67 (3.14)	4.42 (3.24)						
		[0.853]	[0.744]	[1.000]	[0.773]	[0.117]						
HH head is male (indic.)	0.85 (0.36)	0.85 (0.34)	0.87 (0.38)	0.82 (0.35)	0.85 (0.38)	0.82 (0.38)						
		[0.877]	[0.596]	[0.297]	[0.497]	[0.0958]						
HH head age (yrs.)	45.82 (14.09)	46.43 (13.76)	46.6 (14.19)	46.18 (13.90)	46.43 (13.68)	45.97 (13.94)						
		[0.711]	[0.634]	[0.636]	[0.416]	[0.515]						
HH head is literate (indic.)	0.79 (0.41)	0.76 (0.43)	0.74 (0.44)	0.77 (0.42)	0.76 (0.43)	0.73 (0.45)						
		[0.324]	[0.0505]	[0.312]	[0.266]	[0.0278]						
Area farmed (ha.)	3.39 (3.09)	3.31 (3.94)	3.68 (3.96)	3.76 (4.72)	3.78 (3.74)	3.47 (3.20)	1.29 (0.60)	1.24 (0.61)	1.32 (0.64)	1.3 (0.66)	1.35 (0.63)	1.3 (0.61)
		[0.991]	[0.206]	[0.119]	[0.307]	[0.610]		[0.435]	[0.579]	[0.665]	[0.620]	[0.959]
Fertilizer used, maize (kg.)	29.28 (84.11)	23.47 (59.32)	22.28 (95.69)	26 (66.77)	28.55 (167.25)	19.1 (64.90)	1.03 (1.96)	0.93 (1.86)	0.72 (1.69)	0.97 (1.90)	0.68 (1.71)	0.74 (1.68)
		[0.389]	[0.639]	[0.936]	[0.802]	[0.397]		[0.563]	[0.543]	[0.452]	[0.432]	[0.545]
Fertilizer used, maize (kg./ha.)	15.49 (47.45)	12.88 (40.08)	9.93 (48.12)	16.04 (50.52)	11.16 (42.90)	12.55 (61.18)	0.85 (1.65)	0.75 (1.55)	0.57 (1.37)	0.81 (1.64)	0.55 (1.40)	0.61 (1.44)
		[0.150]	[0.683]	[0.340]	[0.933]	[0.757]		[0.412]	[0.514]	[0.400]	[0.472]	[0.725]
Fertilizer used, maize (indic.)	0.22 (0.42)	0.21 (0.41)	0.16 (0.37)	0.21 (0.41)	0.14 (0.35)	0.17 (0.38)						
		[0.728]	[0.693]	[0.370]	[0.433]	[0.795]						
Fertilizer used, all crops (kg.)	80.98 (185.48)	78.51 (166.07)	45.78 (160.67)	61.32 (131.59)	97.13 (378.38)	44.68 (79.48)	2.14 (2.42)	2.2 (2.40)	1.38 (2.12)	1.81 (2.31)	2.13 (2.41)	1.86 (2.21)
		[0.780]	[0.0445]	[0.219]	[0.696]	[0.0102]		[0.778]	[0.0215]	[0.636]	[0.928]	[0.265]
Fertilizer used, all crops (indic.)	0.47 (0.50)	0.49 (0.50)	0.33 (0.47)	0.41 (0.49)	0.48 (0.50)	0.44 (0.50)						
		[0.636]	[0.0340]	[0.763]	[0.889]	[0.658]						
Improved seeds used (kg.)	21.42 (35.89)	26.05 (67.76)	24.14 (42.88)	25.13 (40.15)	30.63 (45.88)	23.05 (42.28)	1.74 (1.82)	1.88 (1.81)	1.75 (1.88)	1.84 (1.89)	2.06 (1.93)	1.7 (1.88)
		[0.282]	[0.289]	[0.167]	[0.118]	[0.966]		[0.358]	[0.513]	[0.255]	[0.258]	[0.460]
Improved seeds used (kg./ha.)	9.82 (19.33)	10.04 (17.00)	8.37 (12.99)	10.42 (16.14)	9.79 (13.77)	9.92 (18.94)	1.37 (1.45)	1.47 (1.43)	1.3 (1.42)	1.43 (1.49)	1.52 (1.43)	1.28 (1.48)
		[0.840]	[0.867]	[0.293]	[0.937]	[0.800]		[0.434]	[0.871]	[0.267]	[0.497]	[0.233]
Maize production (kg.)	2245.29 (2638.79)	2618.03 (8472.55)	2535.92 (4377.18)	2367.6 (3416.12)	2568 (4393.18)	1954.71 (2343.65)	7.26 (0.95)	7.18 (1.01)	7.16 (1.17)	7.15 (1.22)	7.31 (1.04)	7.08 (1.06)
		[0.471]	[0.263]	[0.437]	[0.429]	[0.596]		[0.456]	[0.633]	[0.648]	[0.997]	[0.0721]
Maize yield (kg./ha.)	1059.01 (1586.25)	936.84 (1170.43)	963.51 (1768.79)	937.15 (1122.79)	843.41 (822.88)	899.49 (1927.26)	6.41 (1.01)	6.4 (0.93)	6.29 (1.07)	6.31 (1.13)	6.39 (0.88)	6.24 (1.08)
		[0.200]	[0.577]	[0.376]	[0.0541]	[0.370]		[0.778]	[0.359]	[0.450]	[0.709]	[0.0676]
Maize sold (kg.)	498.67 (1433.67)	647.08 (1888.58)	668.01 (2009.98)	631.06 (1735.13)	599.37 (1247.07)	548.66 (1561.05)	2.97 (3.22)	2.91 (3.27)	2.64 (3.32)	2.87 (3.29)	3.1 (3.33)	2.78 (3.23)
		[0.242]	[0.339]	[0.335]	[0.876]	[0.877]		[0.939]	[0.512]	[0.894]	[0.583]	[0.163]
Maize sold (indic.)	0.5 (0.50)	0.48 (0.50)	0.42 (0.50)	0.47 (0.50)	0.5 (0.50)	0.46 (0.50)						
		[0.844]	[0.247]	[0.891]	[0.398]	[0.111]						
Has irrigation (indic.)	0.05 (0.21)	0.05 (0.22)	0.02 (0.13)	0.02 (0.15)	0.04 (0.20)	0.02 (0.13)						
		[0.660]	[0.231]	[0.491]	[0.751]	[0.187]						
Fertilizer experience (years out of last 9 years)	1.05 (2.19)	1 (2.13)	0.62 (1.57)	0.75 (1.69)	0.68 (1.64)	0.67 (1.73)	0.41 (0.68)	0.4 (0.66)	0.27 (0.55)	0.33 (0.58)	0.29 (0.57)	0.29 (0.56)
		[0.798]	[0.245]	[0.674]	[0.295]	[0.264]		[0.857]	[0.300]	[0.919]	[0.391]	[0.298]
N	258	238	269	296	236	237	258	238	269	296	236	237

Note: Means presented in top row for each variable, with standard deviations in parentheses. Treatments are as described in Table 1. Data are from April 2011 survey, prior to info and match treatments but after voucher treatment. All variables are either time-invariant (head's education, gender, age, and literacy) or refer to season preceding voucher treatment (retrospective reports on 2009-10 season). Section A includes shows all variables in levels, and Section B specifies production variables in log form. In brackets: p-values of test of equality of mean in a given treatment group with mean in pure control group, after partialling-out fixed effects for 32 stratification cells (groups of three nearby localities, within which information and match treatments were randomly assigned). Standard errors clustered at level of 94 localities.

**Table 3: Impacts on Consumption and Assets, 2012-13**

<u>Dependent variable:</u>	log(per capita consumption)	log(assets)	Index of consumption and assets
	(1)	(2)	(3)
Subsidy ( $\alpha$ )	0.084 (0.035)**	0.200 (0.115)*	0.179 (0.070)**
Basic savings ( $\beta_b$ )	0.091 (0.039)**	0.190 (0.112)*	0.183 (0.069)**
Basic savings + Subsidy ( $\alpha + \beta_b + \gamma_b$ )	0.037 (0.036)	0.120 (0.098)	0.094 (0.061)
Matched savings ( $\beta_m$ )	0.099 (0.043)**	0.154 (0.102)	0.175 (0.072)**
Matched savings + Subsidy ( $\alpha + \beta_m + \gamma_m$ )	0.088 (0.038)**	0.160 (0.100)	0.162 (0.066)**
N	1,533	1,534	1,534
R-squared	0.13	0.27	0.21
<u>Complementarity between...</u>			
Basic savings & Subsidy ( $\gamma_b$ )	-0.137 (0.053)**	-0.270 (0.150)*	-0.269 (0.092)**
Matched savings & Subsidy ( $\gamma_m$ )	-0.094 (0.063)	-0.193 (0.138)	-0.193 (0.106)*
Mean of level variable in control group (MZN) (std.dev.)	72 (39)	62,515 (85,413)	0.000 (0.802)
<u>P-value of F-statistic :</u>			
All treatment effects equal	0.493	0.923	0.515
All treatment effects zero	0.034	0.519	0.028
Complementarity parameters jointly zero	0.037	0.195	0.016

\* significant at 10%; \*\* significant at 5%; \*\*\* significant at 1%

Note: Standard errors (clustered at level of 94 localities) in parentheses. Consumption and assets are averaged across 2012 and 2013 surveys. Dependent variables in columns 1 and 2 are specified in log form. Index of consumption and assets is the average of log consumption and log assets variables after each has been normalized by mean and standard deviation in the pure control group (T1). Each regression includes fixed effects for stratification cell (groups of three localities). All regressions include control variables. Control variables for production outcomes are expressed in logs. Approx. 27 Mozambican meticaís (MZN) per US dollar during study period. Test of "all treatment effects equal" is test that  $\alpha = \beta_b = \beta_m = \alpha + \beta_b + \gamma_b = \alpha + \beta_m + \gamma_m$ .

**Table 4: Variance-comparison tests**

	<u>Standard deviation of</u> <u>Log consumption</u>
Pure control group ( $\sigma_c$ )	0.453
Subsidy ( $\sigma_v$ )	0.545
Basic savings ( $\sigma_b$ )	0.502
Basic savings + Subsidy ( $\sigma_{bv}$ )	0.480
Matched savings ( $\sigma_m$ )	0.489
Matched savings + Subsidy ( $\sigma_{mv}$ )	0.446
P-value of variance-comparison F-test:	
<u>Panel A: Versus pure control group:</u>	
$\sigma_c = \sigma_v$	0.004
$\sigma_c = \sigma_b$	0.093
$\sigma_c = \sigma_{bv}$	0.334
$\sigma_c = \sigma_m$	0.225
$\sigma_c = \sigma_{mv}$	0.829
<u>Panel B: Versus subsidy alone:</u>	
$\sigma_v = \sigma_b$	0.196
$\sigma_v = \sigma_{bv}$	0.039
$\sigma_v = \sigma_m$	0.098
$\sigma_v = \sigma_{mv}$	0.002
<u>Panel C: Versus Matched savings + Subsidy</u>	
$\sigma_{mv} = \sigma_b$	0.064
$\sigma_{mv} = \sigma_{bv}$	0.245
$\sigma_{mv} = \sigma_m$	0.162

Notes: Variance-comparison F-tests are two-sided. Log consumption is log of average per capita household consumption across 2012 and 2013.

**Table 5: Impacts on Quantiles of Consumption Distribution, 2012-13**

Quantile regressions

Dependent variable: Percentile of distribution of log per capita consumption in household:

	10th (1)	20th (2)	30th (3)	40th (4)	50th (5)	60th (6)	70th (7)	80th (8)	90th (9)
Subsidy only	-0.041 (0.058)	0.018 (0.038)	0.035 (0.046)	0.002 (0.039)	0.016 (0.053)	0.068 (0.054)	0.063 (0.053)	0.094 (0.080)	0.179 (0.069)***
Basic savings only	-0.014 (0.052)	0.003 (0.043)	0.069 (0.053)	0.088 (0.052)*	0.103 (0.053)*	0.085 (0.050)*	0.086 (0.052)	0.094 (0.070)	0.180 (0.067)***
Basic savings + Subsidy	0.005 (0.062)	0.045 (0.038)	0.056 (0.045)	0.047 (0.049)	0.053 (0.047)	0.029 (0.050)	0.017 (0.059)	-0.009 (0.055)	0.116 (0.082)
Matched savings only	0.075 (0.048)	0.071 (0.040)*	0.067 (0.048)	0.074 (0.055)	0.074 (0.054)	0.019 (0.047)	0.045 (0.058)	0.039 (0.063)	0.167 (0.067)**
Matched savings + Subsidy	0.073 (0.052)	0.089 (0.048)*	0.074 (0.052)	0.107 (0.052)**	0.110 (0.048)**	0.057 (0.059)	0.088 (0.053)*	0.115 (0.074)	0.176 (0.064)***
N	1,533	1,533	1,533	1,533	1,533	1,533	1,533	1,533	1,533
R-squared	0.08	0.09	0.1	0.1	0.1	0.11	0.1	0.09	0.09
Consumption in control group at given pctl (MZN)	36.09	43.67	49.90	54.70	61.97	71.34	82.06	94.57	114.05
<i>P-value of F-statistic :</i>									
All treatment effects equal	0.027	0.236	0.899	0.131	0.398	0.644	0.672	0.316	0.919
All treatment effects zero	0.038	0.191	0.657	0.171	0.141	0.552	0.391	0.326	0.006

\* significant at 10%; \*\* significant at 5%; \*\*\* significant at 1%

Note: Standard errors (clustered at level of 94 localities) in parentheses. Consumption is averaged across 2012 and 2013 surveys before taking logs. Each regression includes fixed effects for stratification cell (groups of three localities). All regressions include control variables. Control variables for production outcomes are expressed in logs. Approx. 27 Mozambican meticaais (MZN) per US dollar during study period.

**Table 6: Heterogeneous Effects by Baseline (Apr 2011) Consumption per Capita**Dependent variable: Log per capita consumption in household (2012-13)

	(1)	(2)	(3)
Any treatment	0.070 (0.023)***		
Subsidy only		0.061 (0.029)**	0.061 (0.029)**
Any savings treatment		0.073 (0.024)***	
Basic savings only			0.112 (0.034)***
Basic savings + Subsidy			0.051 (0.030)*
Matched savings only			0.078 (0.034)**
Matched savings + Subsidy			0.057 (0.033)*
Any treatment * Baseline consumption	0.042 (0.023)*		
Subsidy only * Baseline consumption		0.045 (0.042)	0.045 (0.042)
Any savings treatment * Baseline consumption		0.042 (0.024)*	
Basic savings only * Baseline consumption			0.078 (0.030)**
(Basic savings + Subsidy) * Baseline consumption			0.026 (0.033)
Matched savings only * Baseline consumption			0.050 (0.035)
(Matched savings + Subsidy) * Baseline consumption			0.019 (0.034)
Baseline consumption	0.226 (0.020)***	0.226 (0.020)***	0.226 (0.021)***
N	1,531	1,531	1,531
R-squared	0.36	0.36	0.37
<i>P-value of F-statistic :</i>			
All interaction terms equal	n.a.	0.952	0.469
All interaction terms zero	n.a.	0.181	0.163

\* significant at 10%; \*\* significant at 5%; \*\*\* significant at 1%

Note: Standard errors (clustered at level of 94 localities) in parentheses. Dependent variable (log consumption per capita) averaged across 2012 and 2013 surveys. "Baseline consumption" is log consumption per capita in household in April 2011, normalized to have mean 0 and standard deviation 1. Each regression includes fixed effects for stratification cell (groups of three localities). All regressions include control variables. Control variables for production outcomes are expressed in logs. Approx. 27 Mozambican meticals (MZN) per US dollar during study period.

**Table 7: Heterogeneous Effects with respect to Agricultural Shocks**Dependent variable: Log per capita consumption in household (2012-13)

	(1)	(2)	(3)
Any treatment	0.073 (0.038)*		
Subsidy only		0.128 (0.056)**	0.129 (0.056)**
Any savings treatment		0.057 (0.038)	
Basic savings only			0.035 (0.046)
Basic savings + Subsidy			0.029 (0.050)
Matched savings only			0.091 (0.053)*
Matched savings + Subsidy			0.081 (0.047)*
Any treatment * Bad years	0.026 (0.056)		
Subsidy only * Bad years		-0.074 (0.061)	-0.075 (0.061)
Any savings treatment * Bad years		0.051 (0.057)	
Basic savings only * Bad years			0.078 (0.067)
(Basic savings + Subsidy) * Bad years			0.040 (0.073)
Matched savings only * Bad years			0.065 (0.072)
(Matched savings + Subsidy) * Bad years			0.025 (0.063)
Bad years	-0.003 (0.052)	-0.004 (0.051)	-0.005 (0.051)
N	1,356	1,356	1,356
R-squared	0.12	0.13	0.13
<u>P-value of F-statistic :</u>			
All interaction terms equal	n.a.	0.007	0.067
All interaction terms zero	n.a.	0.025	0.116

\* significant at 10%; \*\* significant at 5%; \*\*\* significant at 1%

Note: Standard errors (clustered at level of 94 localities) in parentheses. Dependent variable (log consumption per capita) averaged across 2012 and 2013 surveys. "Bad years" is number of years respondents report as bad for agriculture in 2012-13 (takes values of 0, 1, or 2). Each regression includes fixed effects for stratification cell (groups of three localities). All regressions include control variables. Control variables for production outcomes are expressed in logs. Approx. 27 Mozambican meticals (MZN) per US dollar during study period.

**Table 8: Impacts on Investments, 2012-13**

<u>Dependent variable:</u>	Fertilizer purchases (1)	Education expenses (2)	Agricultural investments				Non-agricultural investments		Addendum:
			Land (3)	Irrigation (4)	Machinery (5)	Other (6)	Property (7)	Other (8)	Crop production (9)
Subsidy ( $\alpha$ )	0.539 (0.260)**	0.012 (0.185)	0.253 (0.234)	-0.099 (0.190)	-0.044 (0.298)	0.209 (0.156)	0.338 (0.179)*	0.085 (0.180)	0.156 (0.080)*
Basic savings ( $\beta_b$ )	0.074 (0.323)	-0.011 (0.184)	-0.121 (0.226)	-0.005 (0.187)	-0.131 (0.255)	0.332 (0.149)**	0.198 (0.174)	-0.072 (0.196)	0.026 (0.083)
Basic savings + Subsidy ( $\alpha+\beta_b+\gamma_b$ )	0.605 (0.261)**	-0.134 (0.146)	-0.108 (0.186)	0.091 (0.180)	-0.210 (0.259)	0.126 (0.132)	0.093 (0.153)	0.055 (0.157)	0.015 (0.082)
Matched savings ( $\beta_m$ )	0.596 (0.321)*	0.016 (0.147)	-0.152 (0.197)	0.244 (0.196)	-0.423 (0.290)	-0.007 (0.143)	0.106 (0.183)	0.257 (0.218)	0.205 (0.088)**
Matched savings + Subsidy ( $\alpha+\beta_m+\gamma_m$ )	0.465 (0.293)	-0.082 (0.158)	-0.035 (0.214)	0.056 (0.192)	-0.073 (0.238)	0.274 (0.135)**	0.019 (0.171)	0.101 (0.187)	0.038 (0.081)
N	1,530	1,534	1,533	1,534	1,532	1,534	1,533	1,532	1,522
R-squared	0.33	0.13	0.05	0.14	0.06	0.05	0.07	0.08	0.33
<u>Complementarity between...</u>									
Basic savings & Subsidy ( $\gamma_b$ )	-0.008 (0.329)	-0.136 (0.266)	-0.240 (0.327)	0.195 (0.259)	-0.035 (0.370)	-0.415 (0.210)*	-0.443 (0.242)*	0.041 (0.252)	-0.168 (0.108)
Matched savings & Subsidy ( $\gamma_m$ )	-0.670 (0.385)*	-0.110 (0.231)	-0.136 (0.297)	-0.089 (0.284)	0.394 (0.408)	0.072 (0.210)	-0.425 (0.250)*	-0.242 (0.254)	-0.323 (0.118)***
Mean of level variable in control group (MZN) (std.dev.)	1,242 (3,320)	989 (1,352)	199 (766)	275 (1,318)	227 (678)	42 (227)	1,024 (12,540)	433 (3,505)	19,180 (31,087)
<u>P-value of F-statistic:</u>									
All treatment effects equal	0.130	0.820	0.384	0.451	0.704	0.210	0.531	0.767	0.062
All treatment effects zero	0.018	0.883	0.499	0.563	0.722	0.108	0.463	0.834	0.060
Complementarity parameters jointly zero	0.123	0.858	0.764	0.557	0.427	0.041	0.132	0.499	0.028

\* significant at 10%; \*\* significant at 5%; \*\*\* significant at 1%

Note: Standard errors (clustered at level of 94 localities) in parentheses. All variables are averaged across 2012 and 2013 surveys. Dependent variables are specified as  $\log(1+x)$ . Each regression includes fixed effects for stratification cell (groups of three localities). All regressions include control variables. Production outcomes that are used as control variables are expressed in logs. Approx. 27 Mozambican meticaís (MZN) per US dollar during study period.

**Table 9: Impacts on Financial Management, 2012-13**

<u>Dependent variable:</u>	Formal savings (1)	Cash not in banks (2)	Durable goods (3)	Livestock (4)	Crop stocks (5)	Loan balance (6)
Subsidy ( $\alpha$ )	0.642 (0.285)**	0.264 (0.175)	0.088 (0.170)	0.400 (0.255)	0.168 (0.100)*	0.191 (0.306)
Basic savings ( $\beta_b$ )	1.298 (0.306)***	0.141 (0.172)	0.071 (0.134)	0.362 (0.242)	0.132 (0.100)	0.398 (0.252)
Basic savings + Subsidy ( $\alpha + \beta_b + \gamma_b$ )	1.317 (0.299)***	-0.045 (0.191)	0.146 (0.114)	0.074 (0.239)	0.114 (0.096)	0.589 (0.245)**
Matched savings ( $\beta_m$ )	1.403 (0.326)***	0.456 (0.191)**	0.169 (0.144)	0.291 (0.218)	-0.036 (0.138)	0.232 (0.253)
Matched savings + Subsidy ( $\alpha + \beta_m + \gamma_m$ )	1.757 (0.342)***	0.234 (0.190)	0.053 (0.147)	0.271 (0.240)	-0.007 (0.099)	0.461 (0.285)
N	1,534	1,520	1,534	1,534	1,534	1,533
R-squared	0.17	0.1	0.25	0.2	0.14	0.08
<u>Complementarity between...</u>						
Basic savings & Subsidy ( $\gamma_b$ )	-0.623 (0.425)	-0.450 (0.236)*	-0.013 (0.198)	-0.687 (0.342)**	-0.185 (0.137)	0.000 (0.413)
Matched savings & Subsidy ( $\gamma_m$ )	-0.288 (0.423)	-0.486 (0.256)*	-0.203 (0.224)	-0.420 (0.312)	-0.139 (0.168)	0.039 (0.427)
Mean of level variable in control group (MZN) (std.dev.)	3,629 (37,690)	3,548 (11,320)	14,285 (39,652)	33,605 (42,837)	7,636 (14,035)	2,612 (10,102)
<u>P-value of F-statistic :</u>						
All treatment effects equal	0.049	0.182	0.856	0.600	0.374	0.605
All treatment effects zero	0.000	0.080	0.772	0.514	0.319	0.170
Complementarity parameters jointly zero	0.345	0.094	0.503	0.138	0.391	0.994

\* significant at 10%; \*\* significant at 5%; \*\*\* significant at 1%

Note: Standard errors (clustered at level of 94 localities) in parentheses. All variables are averaged across 2012 and 2013 surveys. Dependent variables are specified as  $\log(1+x)$ . Each regression includes fixed effects for stratification cell (groups of three localities). All regressions include control variables. Production outcomes that are used as control variables are expressed in logs. Approx. 27 Mozambican meticaís (MZN) per US dollar during study period.



**Table 10: Impacts on Indices of Intertemporal Activities, 2012-13**

<u>Dependent variable:</u>	Investment index	Financial management index	Index of intertemporal activities
	(1)	(2)	(3)
Subsidy ( $\alpha$ )	0.163 (0.088)*	0.293 (0.105)***	0.217 (0.075)***
Basic savings ( $\beta_b$ )	0.037 (0.091)	0.396 (0.118)***	0.190 (0.079)**
Basic savings + Subsidy ( $\alpha + \beta_b + \gamma_b$ )	0.066 (0.081)	0.366 (0.105)***	0.191 (0.071)***
Matched savings ( $\beta_m$ )	0.081 (0.092)	0.414 (0.104)***	0.223 (0.080)***
Matched savings + Subsidy ( $\alpha + \beta_m + \gamma_m$ )	0.092 (0.088)	0.455 (0.102)***	0.244 (0.080)***
N	1,534	1,534	1,534
R-squared	0.24	0.25	0.29
<u>Complementarity between...</u>			
Basic savings & Subsidy ( $\gamma_b$ )	-0.134 (0.108)	-0.322 (0.162)*	-0.216 (0.103)**
Matched savings & Subsidy ( $\gamma_m$ )	-0.152 (0.121)	-0.252 (0.145)*	-0.197 (0.110)*
Mean of level variable in control group (MZN) (std.dev.)	1.859 (1.083)	5.868 (1.421)	3.578 (1.014)
<u>P-value of F-statistic :</u>			
All treatment effects equal	0.821	0.683	0.962
All treatment effects zero	0.571	0.000	0.015
Complementarity parameters jointly zero	0.378	0.096	0.081

\* significant at 10%; \*\* significant at 5%; \*\*\* significant at 1%

Note: Standard errors (clustered at level of 94 localities) in parentheses. Indices are averages of dependent variables in previous tables after each has been normalized by mean and standard deviation in the pure control group (T1). Investment index does not include crop production variable in calculation of the average. Each regression includes fixed effects for stratification cell (groups of three localities). All regressions include control variables. Control variables for production outcomes are expressed in logs. Approx. 27 Mozambican meticaís (MZN) per US dollar during study period.

**Appendix Table 1: Impact of treatments on attrition from follow-up surveys**

	<u>Dependent variable: Attrition from...</u>			
	1st follow-up survey	2nd follow-up survey	3rd follow-up survey	2nd and 3rd follow-up survey
Subsidy	-0.015 (0.025)	0.054 (0.034)	0.01 (0.025)	0.002 (0.018)
Basic savings	-0.006 (0.024)	0.018 (0.025)	-0.023 (0.017)	-0.006 (0.014)
Basic savings + Subsidy	0.006 (0.024)	0.019 (0.027)	-0.006 (0.019)	-0.017 (0.013)
Matched savings	-0.013 (0.027)	0.047 (0.028)*	0.004 (0.021)	0.003 (0.016)
Matched savings + Subsidy	0.009 (0.027)	0.034 (0.027)	-0.015 (0.025)	-0.007 (0.019)
<i>P-value of F-test, joint signif of all treatment coeffs</i>	0.862	0.582	0.356	0.511
Mean dep var, control group	0.094	0.075	0.071	0.034
Observations	1,589	1,589	1,589	1,589
R-squared	0.03	0.03	0.03	0.03

\*\*\* p<0.01, \*\* p<0.05, \* p<0.1

Note: Standard errors (clustered by 94 localities) in parentheses. Dependent variable is an indicator equal to 1 if respondent attrited from given follow-up survey (i.e., attrition is always with respect to initial study participant list). Each regression includes fixed effects for stratification cell (groups of three localities).

**Appendix Table 2: Impact of treatments on assistance to other households**

<u>Dependent variable:</u>	Indicator: any assistance given to other households	Number of times assisted other households
Subsidy	-0.01 (0.045)	0.196 (0.229)
Basic savings	0.007 (0.043)	0.072 (0.183)
Basic savings + Subsidy	-0.032 (0.047)	-0.291 (0.221)
Matched savings	0.019 (0.047)	0.433 (0.278)
Matched savings + Subsidy	0.057 (0.045)	0.299 (0.246)
<i>P-value of F-test, joint signif of all treatment coeffs</i>	0.416	0.205
Mean dep var, control group	0.607	1.923
Observations	1,533	1,533
R-squared	0.11	0.13

\*\*\* p<0.01, \*\* p<0.05, \* p<0.1

Note: Standard errors (clustered by 94 localities) in parentheses. Dependent variables refer to assistance to other households in 2012 and 2013 surveys. Each regression includes fixed effects for stratification cell (groups of three localities). All regressions include control variables.